

Improving Access and Quality in Early Childhood Development Programs

Experimental Evidence from The Gambia

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Abstract

Early childhood experiences lay the foundation for outcomes later in life. Policy makers in developing countries face a dual challenge of promoting access to and quality of early childhood development services, but evidence on how to manage this trade-off is scarce. This paper studies two experiments of early childhood development programs in The Gambia: one increasing access to services, and another improving service quality. In the first experiment, new community-based early childhood development centers were introduced to randomly chosen villages that had no preexisting, structured early childhood development services. In the second experiment, a randomly assigned subset of existing early childhood development centers received intensive provider training. The analysis finds no

evidence that either intervention improved average levels of child development. Exploratory analysis suggests that the first experiment, which increased access to relatively low-quality early childhood development services, led to declines in child development among children from less disadvantaged households. The evidence supports that these households may have been steered away from better quality early childhood settings in their homes. Comparisons of observationally similar children across experiments reveal that existing early childhood development centers increased language skills by 0.4 standard deviation relative to the community-based alternative, reflecting differences in program quality.

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Improving Access and Quality in Early Childhood Development Programs: Experimental Evidence from The Gambia*

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1. Introduction

Early childhood experiences lay the foundation for outcomes later in life. Extensive research has documented links between early life circumstances and adult outcomes (see Currie and Almond 2011 for a review of the U.S. literature; Nores and Barnett 2010 for evidence outside the U.S.; and Tanner, Candland, and Odden 2015 for developing countries). Due to adverse early life circumstances, in developing countries 250 million children under age 5 risk failure of reaching their development potential (Black et al. 2017).

More than 80 million of these children live in Sub-Saharan Africa, representing two-thirds of children in the region. Countries in Sub-Saharan Africa face two main challenges to promoting children's development. The first challenge is lack of access to early childhood development (ECD) services. Enrollment in pre-primary schooling is 22% in Sub-Saharan Africa, compared to 79% in OECD countries (World Bank 2015).

The second challenge is ensuring the quality of existing ECD services. Even where access exists, program quality can be low and uneven due to resource constraints and low skills of ECD providers. This low quality can result in inadequate cognitive stimulation for children, identified as one of four key risk factors hindering childhood development in developing countries (Walker et al. 2007, 2011).¹

This paper evaluates two experiments to improve access and quality in early childhood development services in The Gambia, one of the poorest countries in the world, with an annual per capita income of 427 USD (World Bank 2016). The experiments tested alternative approaches to deliver a new curriculum intended to stimulate development of children aged 3-6 through structured play. In the first experiment, new community-based early childhood development centers were introduced to randomly chosen villages that had no pre-existing structured ECD services. This intervention increases access to new ECD services.

In the second experiment, a randomly assigned subset of existing ECD centers operating as annexes in elementary schools received intensive provider training to implement the new curriculum. A control group of ECD centers (also annexes) received the new curriculum without the provider training. This intervention improves the quality of existing ECD services.

We find no evidence that either experimental intervention improved average levels of fine motor skills or language and hearing, our key child development measures. We also find no evidence of program effects at other points in the outcome distribution.

Despite these results for the full sample, the community-based ECD intervention had differential effects according to several baseline characteristics. Exploratory analysis suggests that children from more advantaged households scored about 0.4 standard deviations lower in language skills when living in a community randomly assigned to the community-based ECD treatment. These households may have been steered away from better quality early childhood settings for their children, possibly in their homes.² Nevertheless, parents in treatment communities still send their children to these facilities either because they have an erroneous perception of their quality, or because of the value of the time freed up by sending their children to preschool instead of caring for them at home. In contrast, we find no differential effects of the provider training treatment in ECD Annex centers according to children's background.

¹ The other risk factors are stunting, iodine deficiency, and iron deficiency anemia.

² Negative impacts of low-quality ECD have been found elsewhere (see, e.g., S. Berlinski and Schady 2015).

It is noteworthy that the type of centers in each experiment, as well as the nature of the interventions, are very different. Community-based centers are provided when no ECD services exist in the community. There are few or no physical or human resources to build on, which means that the services provided are of limited quality. New but very basic infrastructure needs to be built to house the centers. Teachers are residents in the community with few qualifications and experience, who undergo minimal training to operate these centers. In the type of communities where no centers exist, these community centers are made possible by leveraging minimal investments and local human resources. Currently, they are the only center-based alternative that the government can provide at scale.

In contrast, ECD Annexes operate in the existing elementary school infrastructure. In addition, they employ individuals with much higher qualifications and teaching experience. Like community centers, they also leverage local physical and human resources to provide ECD services, but these resources are of higher quality in communities with an existing elementary school. It is likely that ECD Annexes are of higher quality than ECD community centers. In fact, our own non-experimental evidence suggests that is precisely the case.

The first intervention therefore improves access to ECD center-based services, from a situation of no access whatsoever to one where minimal quality ECD centers are provided. In contrast, the second intervention aims to increase the quality of already existing ECD centers in school annexes, which are of higher quality than ECD community centers even in the absence of treatment.

In sum, we find no evidence that the minimal quality centers that are provided in these settings are helpful for child development. If anything, the opposite is true, especially for children in good home environments. Our findings of no effects for provider training in existing ECD centers suggest that any quality improvements were insufficient to improve child outcomes.

A central aspect of our paper is that, in both experiments, we consider public child care provision which can be delivered at scale, with the available physical and human resources in very challenging settings (as opposed to, for example, small pilot programs implemented by very specialized and highly qualified staff). The ECD centers we study are currently provided by the government in The Gambia at national scale.³ They are similar to the types of ECD centers that can be delivered by the poorest governments in Sub-Saharan Africa.

We contribute to two strands of the literature on pre-school interventions (i.e., center-based programs for 3-6 year olds) in developing countries (for a review, see J. Behrman, Fernald, and Engle 2013). The first strand, corresponding to our evaluation of community-based ECD centers, compares children exposed to an ECD program with children exposed for shorter durations or to home-based care (J. R. Behrman, Cheng, and Todd 2004 for Bolivia; Aboud 2006 for Bangladesh; Berlinski and Galiani 2007, Berlinski, Galiani, and Gertler 2009 for Argentina; Berlinski, Galiani, and Manacorda 2008 for Uruguay; José Rosero and Oosterbeek 2011, Jose Rosero 2012 for Ecuador; Rao et al. 2012, Bouguen et al. 2014 for Cambodia; Martinez, Naudeau, and Pereira 2013 for Mozambique; Bernal and Fernández 2013 for Colombia; Krafft 2015 for the Arab Republic of Egypt; Attanasio et al, 2018, and Pinto, Santos, and Guimarães 2016 for Brazil; Bastos, Bottan, and Cristia 2017 for Guatemala). Of these, only Martinez, Naudeau, and Pereira (2013), Bouguen et al. (2014) and Attanasio et al (2018)

³ The experiments we analyze covered only two of six regions, however. We provide more details on program design in the following section.

are randomized control trials, and only Martinez, Naudeau, and Pereira (2013) study a program in Sub-Saharan Africa.⁴

Most of the aforementioned studies find positive effects on children's development and subsequent school performance, with some exceptions. Consistent with evidence from developed countries, the counterfactual environment for children in ECD programs is often decisive for impact estimates. Where the alternative to ECD programs is a lower quality home environment, estimated effects will be positive. Alternately, when ECD participation substitutes for other forms of schooling, such as early primary school attendance (Bouguen et al. 2014), or reduces parenting quality by increasing maternal labor force attachment (José Rosero and Oosterbeek 2011; Jose Rosero 2012), impact estimates can be zero or negative. These impacts can also vary within a program according to household characteristics (Pinto, Santos, and Guimarães 2016).

A second, smaller strand of the literature evaluates attempts to upgrade the quality of existing ECD programs via teacher training, as in the ECD Annex experiment in this paper.⁵ Despite a few studies using credible identification strategies (Baker-Henningham et al. 2012 for Jamaica, Yoshikawa et al. 2015 and Bowne, Yoshikawa, and Snow 2016 for Chile, Bernal 2015 for Colombia, and Araujo et al 2018 for Peru), there is scant evidence on provider training in Sub-Saharan Africa. Ozler et al. (2016) find that ECD providers in Malawi who were randomly assigned to receive in-service training improved classroom organization and teaching quality, but these behavioral changes increased child development only when paired with parent education. Wolf, Aber, and Behrman (2017) evaluate an experiment with a similar design in Ghana but find gains in child development only when provider training was *not* paired with parent education. These contrasting findings suggest that program content and implementation can be crucial for efforts to improve ECD program quality. Additionally, even when programs to upgrade quality are effective in RCTs, gains may not be sustained when taken to scale. We build on the important contributions of these studies, adding to the thin evidence base on ECD provider training in Sub-Saharan Africa. Importantly, in our context the programs are implemented at scale by public providers, providing evidence that is immediately relevant for policy.

A related literature examines alternative ways of providing early childhood services at scale in poor settings, beyond center-based care, which is often quite expensive. Popular alternatives that have been investigated include home visits, or group-based parenting. Studies from Jamaica (Walker et al. 2006) and Colombia (Attanasio et al. 2014) suggest that these interventions can be quite successful, even when delivered at scale, as in the Reach Up program (<http://www.reachupandlearn.com/>), although these programs have traditionally targeted younger children. It is possible that alternative programs such as these can be more effectively delivered at scale than center-based care programs such as those studied here.

In the next section, we describe the program and data. Section 3 describes our empirical methodology. Section 4 presents experimental results. Section 5 concludes.

⁴ Several other studies have evaluated pre-school attendance in Sub-Saharan Africa (Taiwo and Tyolo 2002 for Botswana; Mwaura, Sylva, and Malmberg 2008; Malmberg, Mwaura, and Sylva 2011 for Kenya, Tanzania, and Uganda; Zuilkowski et al. 2012 for Zambia), but these studies lack credible identification strategies to distinguish the role of ECD exposure from unobserved child or parent attributes.

⁵ Studies of efforts other than provider training intended to upgrade ECD quality in developing countries include Armeccin et al. (2006) for the Philippines; He, Linden, and MacLeod (2009) for India; and the many follow-ups to the Mauritius Child Health Project (Raine, Venables, and Mednick 1997; Raine et al. 2001, 2003, 2010).

2. Program Description and Data

The Gambia is a small West African country with population 1.9 million and per capita annual income of 427 USD (World Bank 2016), making it one of the 10 poorest countries in the world (World Bank 2015). Its education system is divided into six numbered administrative regions. Region 1 is the capital, Banjul, on the Atlantic coast, with Regions 2-6 located in increasingly remote areas toward the East. Regions become more rural and poorer as distance from the capital increases.

Enrollment in ECD programs among Gambian children aged 3-6 was 22% in 2007 (Zoyem 2010). The Gambian government has sought to increase this proportion, with a goal of integrating ECD programs into the standard primary school sequence. Other than informal home care, ECD services exist in three forms in The Gambia: 1) private centers, located mostly in relatively urban areas and serving richer children; 2) public centers which are built as annexes to primary schools (hereafter referred to as ECD Annexes); and 3) community-based centers, which are publicly run, stand-alone facilities located in communities without primary schools. ECD Annexes and community-based centers do not charge fees.

To increase ECD access and improve quality, the Gambian Ministry of Basic and Secondary Education (MoBSE, hereafter “the ministry”) implemented a new early childhood curriculum in 2012. The curriculum, known as Gambia Open and Active Learning Spaces (GOALS), aims to promote creativity, problem solving, and confidence through structured play. Each day includes time for group activities, games to promote critical thinking, physical development, and music/singing. The curriculum runs for 40 weeks annually, concurrent with the academic calendar. Activities run for four hours daily Monday through Thursday and three hours on Friday. Teachers are expected to spend an additional 1.5 hours each day to prepare for the next day’s session. Additionally, a monthly meeting is held with parents to discuss the program and children’s progress.

The curriculum was deployed at all ECD Annexes and community-based centers nationally. To assess the effectiveness of different approaches to curricular delivery, the ministry experimentally implemented two interventions in parallel.⁶ Both experiments occurred in Regions 2 and 6, with treatments assigned at the village level. One intervention aimed to increase access to ECD center-based care, while the other aimed to improve the quality of existing center-based care.

In the first experiment, the ministry built new community-based centers in randomly chosen villages that had no pre-existing structured ECD services. Each community-based center delivered the new GOALS curriculum. Management committees, comprised of parents and community leaders, were formed to oversee the centers. We refer to this as the community-based treatment, and the corresponding control group of eligible villages without ECD services as the pure control.

In the second experiment, a teacher training program was delivered to a randomly chosen subset of existing ECD Annex centers. Teachers received intensive training in the new curriculum in three sessions between September 2012 and September 2013. The trainings lasted five, eight, and eight days, respectively. The Gambia office of ChildFund, an international NGO, conducted the trainings, which were financed by the Japan Social Development Fund and World Bank. We refer to the experimental groups in this case as ECD Annex treated and ECD Annex control.

Teachers in the community-based ECD centers of the first experiment also received the training provided to the ECD Annex treatment group. Compared to the pure control group, the community-

⁶ A third experiment, on a different ECD program, was conducted for children ages 0-3 (Blimpo, Carneiro, Jervis and Pugatch, 2018), who were too young to be eligible for the services studied in this paper. This experiment occurred in a separate set of communities than the experiments described here.

based treatment therefore represents an increase in access to new ECD services, including site construction, formation of management committees, curriculum, and teacher training. All ECD Annexes in the program regions received the new curriculum, but only the ECD Annex treatment group received the teacher training. The ECD Annex treatment therefore represents an attempt to improve quality through teacher training, since the same curriculum was delivered to all annexes. Table 1 summarizes the research design.

Table 1: Experimental design

Experiment	Sample	Control	Treatment
1	Villages without pre-existing ECD services, Regions 2/6	Pure control <ul style="list-style-type: none"> No structured ECD services 	Community-based <ul style="list-style-type: none"> Construction of ECD center GOALS curriculum Teacher training Management committee
2	Villages with ECD Annex centers, Regions 2/6	ECD Annex control <ul style="list-style-type: none"> GOALS curriculum 	ECD Annex treatment <ul style="list-style-type: none"> GOALS curriculum Teacher training

Figure 1 shows a map of sites included in the sample. The bulk of the sample sites are in Region 6, one of the poorest and the most remote regions of the country.⁷ Treatment was stratified by region to ensure sample balance. We include a dummy variable for Region 2 to account for this stratification throughout all analyses. The sample for the ECD Annex experiment includes 26 treatment sites and 27 controls, while the sample for the community-based ECD experiment includes 40 treatment sites and 51 controls.⁸

The new curriculum and initial teacher training began in September 2012, the start of the 2012-2013 academic year. A baseline survey was conducted before the beginning of the academic year, in May-July 2012. A representative sample of 16 households with children in the relevant age group (3-6 years) was taken from each community eligible for treatment. In households with multiple eligible children, one was randomly sampled. The household head and main caregiver for the eligible child

⁷ Table SA1 of the Supplemental Appendix uses the 2003 Census, the most recent conducted before the program began, to compare communities in program Regions 2 and 6 with the rest of the country (Regions 1 and 3-5). Regions 2 and 6 differ from each other on many dimensions, with Region 2 more populous, educated, and developed. These differences reflect the more urban character of Region 2 compared with rural and remote Region 6. Stratifying the sample by region ensures that these differences are not spuriously correlated with treatment. Regions 2 and 6 also differ from the rest of the country, as shown by the many significant differences reported in columns (4)-(5) in the table. These differences suggest that the treatment effects reported in this study may not generalize to all regions of the country. However, the heterogeneity between Regions 2 and 6 ensures that the program occurred in a broad range of contexts found within The Gambia. The supplemental appendix is available upon request.

⁸ The sample of ECD Annexes is smaller than the community-based ECD experiment because the latter forms an exhaustive list of ECDs in the two regions. A larger number of control sites were sampled in the community-based ECD experiment because these were also used as control sites for a separate experiment with children ages 0-3 (M. P. Blimpo et al. 2018), allowing for economies of scale in data collection. We dropped one ECD Annex treatment site because none of the sampled children met the age eligibility criteria according to their birth certificates. There are two fewer community-based ECD and pure control sites in the baseline data than in the full sample for the same reason.

were interviewed, with modules on household assets, expenditures, employment, demand for ECD services, health (own and the child's), parenting activities, and attitudes towards disciplining the child. Anthropometric measurements (height, weight, brachial circumference) were taken from children.

An endline survey was conducted in November-December 2013, 14 to 15 months after treatment began. The endline survey asked similar questions as the baseline, with an additional module to assess caregiver knowledge of childhood nutrition and health. Endline participants included baseline households and newly randomly sampled households, allowing for an increase in the sample size of analyses using only endline data. We analyze attrition from the survey in the Data section. Figure 2 shows a timeline of the project and research milestones.

In addition to the survey modules previously mentioned, children in sampled households were given the Malawi Developmental Assessment Tool (MDAT), a test of child development designed for rural Africa (Gladstone et al. 2010). The test consists of two modules: 1) fine motor skills, and 2) language and hearing.⁹ Each module has multiple versions tailored to different age ranges. The fine motor skills tool asks children to complete tasks such as stacking blocks in various configurations, placing pegs in a board, and determining the relative weight of objects. The language and hearing tests require children to point to body parts by name, identify the names or uses of objects, identify the letters in one's name, and similar tasks. Each item on the tests is marked as complete or incomplete, with the overall score determined by the total number of completed items. Enumerators received five days of training on the tests before administering them in the field. Due to interviewer time constraints, enumerators were instructed to administer the test to children in a random subsample of surveyed households.

Appendix A presents selected questions from the MDAT versions used in the surveys. The baseline and endline used the MDAT versions intended for children aged 36-59 and 53-76 months, respectively (though children completed the tests regardless of their age at the time of the surveys). We assessed the reliability of the MDAT in our sample by calculating Cronbach's α separately by module (fine motor and language and hearing) and survey wave (baseline and endline), using all available MDAT scores in the survey (including children from the pure control group and community-based ECD treatment). The Cronbach α values ranged from .82-.88, indicating high reliability to measure the underlying constructs. We also check internal consistency of MDAT scores in our sample by examining their relationship with child age and household wealth. Figure 3(a) shows that both fine motor and language skills progress approximately linearly with age. Figure 3(b) plots MDAT scores against deciles of a wealth index, with scores generally increasing in wealth, particularly for the top deciles. The figure also shows the progression of height-for-age with wealth as a comparison.¹⁰

Program implementation, though largely successful, encountered significant challenges. Fourteen of 40 community-based ECD treatment sites reported implementation problems. The most common problems reported were absent or sick teachers or lack of materials. Two sites reported that their facility had not been constructed, forcing teachers to provide services outdoors. Because of administrative issues, many teachers went several months without receiving their stipends, leading to

⁹ Assessments were conducted in the mother tongue or English, whichever was more comfortable for the child.

¹⁰ Figure 3(b) plots our preferred measure of MDAT scores used throughout the paper. We adjust MDAT scores for age by regressing a child's raw score on age (in months), age squared, and a female dummy, then dividing the residual by its standard deviation. We run this procedure separately for the fine motor and language and hearing modules and for baseline and endline. We calculate height-for-age using the international benchmarks defined by the World Health Organization. The wealth index is the first principal component of household asset ownership. Figure 3(a) does not adjust MDAT scores by age.

some resignations and replacements. No ECD Annex treatment sites reported implementation problems. We provide additional details on implementation in Section 4.1.

3. Methodology

We can analyze the effect of each experimental treatment through a comparison of mean outcomes between children in treated and control communities. We cluster all standard errors at the community level to adjust for correlated outcomes among units exposed to the same treatment. To test for differences between treatment and control groups, we further adjust for the stratification of treatment status within regions. We make this adjustment by regressing the outcome on an indicator for treatment and a dummy for whether the community is in Region 2:

$$y_{ic} = \alpha + \beta D_c + \gamma Region2_c + \varepsilon_{ic} \quad (1)$$

where i indexes the child, c indexes the community, y is the outcome, D and $Region2$ are indicators for treatment and Region 2, respectively, and ε is an error term. We estimate this equation separately for the community-based ECD and ECD Annex experiments, with the definition of treatment changing accordingly. In each case, the coefficient of interest is β , which measures the difference in mean outcomes between children exposed to each treatment compared to the corresponding control group. We will also run versions of equation (1) that include the baseline outcome y_{0ic} on the right-hand side in order to improve the precision of our estimates.

Our main outcomes of interest are scores for the two MDAT modules, fine motor skills and language and hearing. We normalize each score by child gender and age, using children from both experiments as the underlying population (outcomes are then measured in standard deviations). Additionally, we analyze subsets of MDAT items which are most closely aligned with the structured play format of the GOALS curriculum, such as stacking and counting blocks. For each subset, we measure the percentage of items successfully completed by the child.

The parameters estimated by equation (2) are intent to treat estimates. For the ECD community center experiment they are likely to be a good approximation to the average treatment effect since there are little to no reports of eligible households refusing enrollment in these centers (as we discuss below, this comes from indirect reports from program staff, since we have not collected this information in our surveys). For the ECD Annex experiment, we do not expect the treatment to affect enrollment rates, which are already very high, bordering 90%. Therefore, a good approximation to the average treatment effect of teacher training for children attending ECD Annexes is (approximately) $\beta/0.9$.

4. Experimental results

4.1. Attrition, baseline balance, and implementation

Table 2 presents sample sizes and analyzes attrition. The sample includes all children aged 3-6 at the time of the baseline. Panel A, columns (1)-(2) show the number of children sampled in baseline and endline, respectively, separately by experiment and treatment group. The endline sample is split

into two categories: those who appeared in the original (baseline) sample, and those newly sampled to increase the sample size. Columns (3)-(4) of Panel A present the same information but restrict the sample to those who completed the MDAT fine motor skills and language and hearing tests. Some children present in the baseline survey completed the MDAT at the endline but not the baseline, and therefore appear in the “original sample” group of columns (4) although they lack a baseline score. For this reason, in column (5) we show the number of children who completed the MDAT in both baseline and endline.

This table shows that, for both experiments, only about half the children in each survey wave are administered the test. This happened because there were not enough funds for testing children in all households. The set of households where testing was conducted was selected at random in both survey waves.

Nevertheless, there are some imbalances in the characteristics of test takers and non-test takers, which occur purely by chance. At baseline, test completers come from statistically significantly larger households, are more likely to have a household head in agriculture and have higher vaccination rates than non-completers in the community-based ECD experiment. In the ECD Annex experiment, test completers are younger, have more educated parents, and higher vaccination rates than non-completers. These differences should be kept in mind when interpreting our results, which apply only to children completing the test. Fortunately, test completion is uncorrelated with treatment in either experiment, increasing confidence in the internal validity of our estimates.¹¹ We later check robustness of results when using multiple imputation to infer the outcomes of children who did not take the MDAT.

Panel B of Table 2 analyzes sample attrition. We define two types of attrition: 1) attrition from the survey, in which a household that completes the baseline survey fails to complete the endline survey, and 2) attrition from the test, in which a child who completes the baseline MDAT does not complete the endline MDAT. Columns (1)-(3) of Panel B show the sample proportion of each type of attrition by treatment status. Attrition in the community-based ECD experiment exceeds that of the ECD Annex experiment. Within each experiment, however, neither type of attrition differs significantly between treatment and control groups, as shown by the p-values in column (5). In what follows, we restrict attention to children who complete the MDAT, unless otherwise indicated.

Table 3 presents baseline characteristics and tests for balance between treatment and control groups within each experiment. Columns (1)-(4) show the control group mean, treatment group mean, difference, and corresponding p-value for the community-based ECD experiment. Columns (5)-(8) repeat the exercise for the ECD Annex experiment. Baseline values across all groups demonstrate the economic and social disadvantages faced by children in The Gambia. Height-for-age, which proxies for early life nutrition, is more than 1.25 standard deviations below the international average for all groups. Average completed schooling of mothers is 2 years or less. Mean household expenditure per capita ranges from US\$418-523. And 13-18% of children were ill at the time of the survey, and mothers report high levels of mental distress (measured as the percentage of indicators of distress, such as feeling lonely, sad, or fearful, experienced “most of the time”).

Looking across experiments, children in the community-based ECD experiment score lower in fine motor skills, language and hearing, and height-for-age than those in the ECD Annex experiment. These disadvantages are not surprising, as the presence of ECD Annexes in these communities suggests greater opportunities for investment in children. Notably, household socioeconomic indicators are not uniformly higher in the ECD Annex experiment communities. Children in the ECD community

¹¹ We report these results in Tables SA2-SA3 of the Supplemental Appendix available upon request.

experiment are younger than those in the ECD Annex experiment, reflecting the relative ages of children who tend to enroll in each type of facility. Because many Gambian children walk to schools unaccompanied by adults, parents are reluctant to send younger children to schools. ECD Annexes tend to be farther away than the village-based community centers, and therefore attract older children.

Within experiments, treatment and control groups are broadly similar. Fine motor skills, language and hearing, and height-for-age are not significantly different between treatment and control in either experiment. Differences in most other characteristics are also not statistically significant, with a few exceptions. Children in the control group of the community-based ECD experiment are 14 percentage points more likely than the treatment group to be attending an ECD program at baseline, significant at 1%. There are also a few differences that are significant at the 10% level, including a greater proportion of treated children with a household head in agriculture and lower assets. In the ECD Annex experiment, treated children have lower household expenditure than control children, significant at 10%. Below we discuss results controlling and not controlling for these variables.

Despite these imbalances, we have no reason to believe that randomization was compromised. The randomization procedure was carefully monitored by the research team and compliance with randomization was perfect. Instead, we attribute baseline differences to chance. Controlling for baseline MDAT scores, which summarize the cumulative effect of early childhood investments, should mitigate spurious treatment effects estimates. Additionally, we estimated treatment effects controlling for the variables imbalanced at baseline and find similar results.¹²

Table 4 reports various measures of attendance at ECD centers and program implementation among the treated communities within each experiment. The data are from a monitoring survey of ECD center administrators and teachers taken in 2013, midway through the program. The top rows of the table show that ECD centers were well attended in both experiments, with an average of more than 45 children registered per center and more than 30 children present on the day of the monitoring visit. Attendance rates on the day of the visit were significantly higher in community-based ECD centers than in treated ECD Annexes, reflecting strong demand in communities that previously lacked structured ECD services. The pupil-teacher ratio exceeds 30 in both experiments, limiting the quality of services that teachers can provide.

Take up rates of ECD services were not recorded. They are however quite high. In ECD community villages there are no reports of households refusing the services nor of waitlists, so take-up rates should be taken to be 100%. Regarding ECD Annex villages (treatment and control), Table 3 reports that baseline rates of prior ECD attendance are 86 to 90%. Given that the main public provider of ECD center-based services in these villages are these ECD Annexes, these are good estimates for take-up rates in the communities in which the school is located.¹³

The bottom rows of Table 4 report teacher characteristics. Compared to treated ECD Annexes, teachers in community-based ECD centers are significantly younger, less likely to have completed secondary school, and have less teaching and ECD experience. They are also more likely to be a village resident. In other words, the modal provider in community-based ECD centers is a local volunteer without previous experience, compared to the professional staff in ECD Annexes. These differences also have implications for the expected quality of service delivery within each experiment, a subject to which we return below. Nonetheless, survey enumerators report higher levels of student engagement in community-based ECD centers than in ECD Annexes (bottom rows of table), though only the

¹² We report those results in Table SA4 of the Supplemental Appendix available upon request.

¹³ The take-up rate of the ECD Annexes is likely lower for more distant communities within the school's catchment area.

“children taking initiative” category differs significantly. These results suggest that teacher effort in community-based ECD centers was at least as high, if not higher, than in ECD Annexes.

4.2. Endline outcomes

We report endline MDAT scores and treatment effects estimates in Table 5. The first two columns show the control and treatment mean, respectively, for the community-based ECD experiment. Columns (3) and (4) show estimates of equation (1), with and without controlling for the baseline outcome. Columns (5)-(8) present the same information for the ECD Annex experiment. In the community-based ECD experiment, the treatment group scores below the control in both language and hearing and fine motor skills. Although this pattern is surprising, neither difference is statistically distinguishable from zero, regardless of whether we control for the baseline score. We also fail to find significant treatment effects when disaggregating each MDAT module by subsets of items most closely associated with the ECD curriculum. Overall, we find no evidence of differences in average scores based on exposure to community-based ECD services.

When dropping sites that reported implementation issues, the treatment effect of community-based ECD on language and hearing falls to -0.24 standard deviation, significant at 10%. Including the baseline score increases the magnitude to -0.29 standard deviation, significant at 5%. The result is surprising, since we expect estimated benefits of community-based ECD to increase when omitting sites with implementation issues. A potential explanation is reporting bias: more conscientious administrators are more likely to report implementation problems, so excluding them leaves only the worst managed sites in the sample. We find no significant differences in fine motor skills when dropping sites with implementation issues. We omit results for brevity but present them in the Supplemental Appendix.¹⁴

In the ECD Annex experiment, the treatment group scores higher on both language and hearing and fine motor skills, but the differences are not precisely estimated. Among subsets of items, the only significant differences are for counting and ordering rows of items, though here the treatment group scores below the control, by 9 and 8 percentage points, respectively.

The mean effects presented in Table 5 might mask changes in other features of the outcome distribution between children in the treatment and control groups of each experiment. We explore this possibility in Figure 4, which plots the distributions of MDAT scores. Figure 4(a) shows kernel density estimates for the community-based ECD experiment, with baseline scores plotted in the first row and endline scores in the second row. Figure 4(b) shows the corresponding density estimates for the ECD Annex experiment. Comparing densities within a column shows how the distributions of each MDAT module shift between baseline and endline.¹⁵

¹⁴ See in Table SA5 of the Supplemental Appendix. Supplemental Appendix Table SA6 presents results when using multiple imputation to infer MDAT scores of children who did not take the test. Following Little and Rubin (2014), we use the baseline variables reported in Table 3 to predict MDAT scores of children with missing scores. We impute missing values 50 times, with inference accounting for the imputation procedure. We impute only overall scores, not subsets of test items, because we are less confident in the results at this level of granularity. Our results for overall scores are of the same signs and similar magnitudes as the main results, with the treatment effect for language in the community-based experiment negative and significant at 10%. These results suggest that the high proportion of children with missing MDAT scores reduces statistical power but does not qualitatively change results. The supplemental appendix is available upon request.

¹⁵ Because we rely on different versions of the MDAT in baseline and endline, the densities are informative about relative changes in the treatment and control distributions, but not of absolute changes in child development.

Beginning with fine motor skills for the community-based ECD experiment in the first column of Figure 4(a), the treatment group lies slightly to the left of the control distribution at baseline. At endline, the mode of the treatment group distribution is to the right of the control group mode. The treatment group's thicker left tail and thinner right tail make the overall change unclear, however. For language skills in the second column, again the treatment group distribution lies slightly to the left of the control group, with only minor differences apparent at endline. In short, we fail to find strong evidence of relative shifts in the MDAT score distributions in the community-based ECD experiment.

Turning to the ECD Annex experiment in Figure 4(b), at baseline we see similar central locations for fine motor skills between the treatment and control groups, though the treatment group has a thicker right tail. By the endline, the treatment group distribution has shifted right relative to the control. In language and hearing, the treatment group begins to the right of the control group distribution at baseline, with the difference somewhat more pronounced at endline. These differences are consistent with the positive point estimates for the ECD Annex treatment coefficient in Table 5.

Although suggestive, the preceding visual inspection of unconditional distributions should not substitute for formal analysis. Even formal tests for equality of unconditional distributions would fail to account for the stratification of treatment assignment by region, or for differences in baseline outcomes. To overcome these limitations, we estimate equation (1) using quantile regressions. For each MDAT module, we estimate the coefficient on treatment status for the 5th-95th (conditional) quantiles, in increments of 5. We continue to cluster standard errors by community, the unit of treatment.

Figure 5(a) presents results for the community-based ECD experiment, with results for the ECD Annex experiment in Figure 5(b). In each graph, the thick black line plots our estimates of the quantile treatment effects. The gray shaded area shows the 95% confidence interval around these estimates. For comparison, we also plot the OLS estimate of the average treatment effect in the long-dashed line, while the short-dashed line shows its 95% confidence interval. The red line shows the gridline for zero treatment effect.

For the community-based ECD experiment in Figure 5(a), a downward slope appears in quantile treatment effect estimates of fine motor skills beginning around the 40th percentile. For language and hearing, estimates remain near the OLS coefficient across all quantiles. For the ECD Annex experiment in Figure 5(b), a downward slope appears in quantile treatment effect estimates around the 40th percentile for fine motor skills and the 20th percentile for language and hearing. These downward slopes suggest that treatment benefitted the middle of the outcome distributions most, with smaller benefits at the top of the distribution. Across both experiments and outcomes, however, estimates of quantile treatment effects are noisy. In fact, confidence intervals include zero around all the quantile effects estimated, suggesting that any differences between treatment and control distributions implied by Figure 4 are spurious. Results are similar when augmenting the quantile regression specification with baseline outcomes, or when using the nonparametric quantile treatment effect estimator of Firpo (2007).¹⁶

4.3 Heterogeneous treatment effects

Although we have found no evidence of significant treatment effects for our main measures of childhood development in the analysis thus far, the null effects for each experimental sample might mask significant effects for subgroups of children. This analysis is more exploratory since, although

¹⁶ Results are reported on Figure SA1 in the supplemental appendix available upon request.

we intended to explore heterogeneous impacts according to child and household characteristics, our study was not explicitly designed to detect them. We examine the following subgroups: male and female; below and above median household assets; whether the child’s mother ever attended school; below or above median quantity of stimulating objects in the home (based on a predetermined list of items); whether the child was sick in the past three days; and below and above median mother mental distress, created from a series of questions about mental health. We chose these groups to analyze whether treatment effects vary by child sex, socioeconomic status, and child and parental health. Subgroups defined as below/above median use both experimental samples to determine the threshold; results are similar when defining the median within each experimental sample.

For each pair of subgroups, we estimate the regression:

$$y_{ic} = \beta_a D_c * \mathbf{1}(\text{group } a)_{ic} + \beta_b D_c * \mathbf{1}(\text{group } b)_{ic} + X_{ic}\theta + \gamma \text{Region}2_c + \varepsilon_{ic} \quad (2)$$

where $D_c * \mathbf{1}(\text{group } a)_{ic}$ is the interaction between treatment and membership in group a (e.g., male), $D_c * \mathbf{1}(\text{group } b)_{ic}$ is the analogous term for group b (e.g., female), and X is a vector of membership dummies for *all* subgroups listed above (not merely those included in the interactions). In other words, we run several versions of equation (2) in order to test for different subgroup interactions, but all control variables remain the same across equations. Our coefficients of interest are β_a and β_b , which measure differential treatment effects for groups a and b relative to those same groups of children in the control group. For instance, when groups a and b are male and female, then β_a and β_b measure treatment effects for male and female children, respectively (the constant is omitted for identification and ease of interpretation).

Within each experiment and child development outcome (fine motor skills and language and hearing), we stack equation (2) across all subgroups and estimate the system jointly. This specification accounts for correlations among error terms in the system and permits hypothesis testing across equations. In addition to testing for significance of all subgroup interaction terms, we are interested in how treatment effects vary by household socioeconomic status (SES). We define low SES as below median household assets, mother did not attend school, and below median stimulating objects in the home, while high SES is the complement of these groups. We then conduct the following joint hypothesis tests:

1. $H1: \beta_{\text{low assets}} = \beta_{\text{mother no school}} = \beta_{\text{low stimulation}} = 0$
2. $H2: \beta_{\text{high assets}} = \beta_{\text{mother attended school}} = \beta_{\text{high stimulation}} = 0$
3. $H3: (\beta_{\text{low assets}} + \beta_{\text{mother no school}} + \beta_{\text{low stimulation}}) = (\beta_{\text{high assets}} + \beta_{\text{mother attended school}} + \beta_{\text{high stimulation}})$

Hypothesis $H1$ tests whether children from low-SES households have non-zero treatment effects. Hypothesis $H2$ is the analogous test for children from high-SES households. Hypothesis $H3$ tests whether treatment effects for low- and high-SES groups differ from each other.

Table 6(a) presents estimates of equation (2) for the community-based ECD experiment, with p-values of hypotheses $H1$ - $H3$ reported at the bottom of the table. We find negative and statistically significant treatment effects for several subgroups: in fine motor skills, for those with above-median stimulating objects at home; and in language, for females, above-median household assets, mother attended school, above-median stimulating objects at home, child not sick in last three days, and both

below- and above-median mother mental distress. The magnitudes range from -0.2 to -0.4 standard deviations. The pattern suggests that females and children from more advantaged households experienced slower language and hearing development via exposure to the community-based ECD program. Our joint hypothesis tests confirm statistically significant negative treatment effects in language for high-SES children. We also find that high-SES children have smaller treatment effects than low-SES children in both fine motor skills and language. Although surprising, these results are plausible if better-off households are able to provide a higher quality home environment than the community-based alternative.¹⁷ A related possibility is that the curriculum was not sufficiently targeted to improving the skills measured by the MDAT, resulting in slower development for children from homes better able to foster these skills.

For the ECD Annex experiment (Table 6(b)), most point estimates are positive, but few subgroups of children have treatment effects significantly different from zero. None of the joint hypothesis tests is statistically significant. We conclude that there is no pattern of treatment effect heterogeneity in the ECD Annex experiment among the groups considered.¹⁸

4.4 Home investments

The final outcome we consider is home investments. Parents of children exposed to either treatment may alter their investment in their children in response. Specifically, the focus of the new GOALS curriculum on stimulation through structured play may encourage parents to increase play activities at home. We therefore explore whether treatment assignment increases the time mothers report playing with their children. It is also possible that parents substitute away from private investments in children in response to higher public investments, in the form of ECD centers. Such a mechanism could even be strong enough generate negative total impacts of these interventions on child outcomes.

We use the full sample of children, regardless of whether they completed the MDAT, in order to maximize statistical power.¹⁹ Table 7, column (1) presents estimates of equation (1), using mother's play time (minutes/day) as the outcome and assignment to community-based ECD as the treatment. The coefficient on treatment is positive, but small (5 minutes, compared to the control mean of 184) and not statistically significant. Columns (2)-(7) present estimates of equation (2), following the same format as Table 6. None of the subgroup treatment effect estimates is significantly different from zero, either individually or jointly. These results suggest that negative treatment effects for some subgroups exposed to community-based ECD are not driven by reductions in home investments.

¹⁷ We find no significant treatment effects for the community-based ECD treatment when splitting the sample by whether the child attended ECD at baseline. These results suggest that the baseline imbalance according to this characteristic is not driving results. Results not shown but available upon request.

¹⁸ We conduct three sets of robustness checks by altering the controls used in equation (2): 1) using no additional controls (i.e., only the Region 2 dummy and main effects of the subgroups tested in each equation, rather than dummies for all subgroups); 2) including indicators for whether the baseline outcome was positive or negative as additional subgroups; and 3) including controls for variables imbalanced at baseline in Table 3. The general pattern of results remains, though some subgroups no longer have statistically significant treatment effects as more controls are added. Importantly, however, we find statistically significantly different treatment effects between high- and low-SES children in the community-based ECD experiment across all specifications. We report results in Tables SA7-SA9 of the Supplemental Appendix. The supplemental appendix is available upon request.

¹⁹ Supplemental Appendix Table SA10 presents robustness checks using the MDAT sample and an additional measure of home investment (mother's total interaction time with children). Results are similar across these alternative specifications. The supplemental appendix is available upon request.

Columns (8)-(14) of Table 7 repeat the exercise for the ECD Annex experiment. Although the overall treatment effect is not statistically distinguishable from zero, we find significant increases in play time for several subgroups. These effects are large, ranging from 36-104 minutes, relative to the control group mean of 162. These increases in home investments make the null findings for the ECD Annex treatment more notable, as the combination of increased teacher training and home-based stimulation failed to improve measures of childhood development.

Table 8, column (1) presents estimates of equation (1), using stimulating objects (measured as the proportion from a predetermined list) as the outcome and assignment to community-based ECD as the treatment. The coefficient on treatment is positive, but small (3 percent points, compared to the control mean of 30 percent) and not statistically significant. Columns (2)-(7) present estimates of equation (2), again following the same format as Table 6. Although treatment effects are positive and significant for male children and households with below-median assets, these are not the groups for whom we found significant treatment effects in Table 6. These results suggest that negative treatment effects for some subgroups exposed to community-based ECD are not driven by reductions in home investments.

Columns (8)-(14) of Table 8 repeat the exercise for the ECD Annex experiment. We find no significant treatment effects for any group, either individually or jointly. These results are consistent with our null findings for the effect of this intervention.

4.5. Estimating the relative effectiveness of ECD Community Centers and ECD Annexes

We now show how we conduct non-experimental comparisons of the effectiveness of ECD Community Centers versus that of ECD Annexes. This is important because our two experiments concern these two potentially very different types of centers. In fact, as we have previously seen, teachers in Annexes have much more experience and education than those in Community Centers. Consistent with this, we document in this section that Annexes are of higher quality than Community Centers, even in the absence of the teacher training treatment.

4.5.1 Method

As discussed in the introduction, meeting the twin goals of access and quality in early childhood development services forces governments to make difficult choices. Comparing the efficacy of the community-based ECD treatment, which attempts to expand access to ECD services, with that of the ECD Annex treatment, which aims to increase quality, matters greatly for policy. In addition, it is important for our paper to establish whether there are important quality differences between these two types of center-based care. However, differences in the populations eligible for the two experiments present challenges for this comparison, as noted earlier.

We attempt to overcome this challenge by comparing *ex ante* similar children across experiments. Specifically, we compare the experimental treatments against each other using an inverse propensity score weighting estimator, which combines matching and reweighting to improve comparability between treated and control units (Hirano, Imbens, and Ridder 2003; Abadie 2005).²⁰ We first match on the propensity score by estimating the probability that the child is in the ECD Annex experiment and not the community-based ECD experiment (regardless of treatment status

²⁰ The exposition in this section closely follows Giordano and Pugatch (2017).

within each experiment).²¹ We then remove children falling outside the common support of the propensity score distributions for each experiment.

Using this trimmed sample, we weight by the inverse propensity score and estimate:

$$y_{ic} = \alpha + \beta_1 community_c + \beta_2 Annex_c + \gamma Region2_c + \varepsilon_{ic} \quad (B.1)$$

where *community* is an indicator for being in the community-based ECD treatment group, *Annex* is an indicator for being in the ECD Annex experiment, and all else is as in equation (1). Because the community-based ECD control group is the omitted category, the coefficients β_1 and β_2 measure the treatment effect of being in each of the corresponding groups relative to the pure control.

The identifying assumption is that treatment assignment is orthogonal to unobserved characteristics that also affect the outcome. In the case of the community-based ECD treatment, this is ensured by random assignment, whereas in the case of the ECD Annex experiment we rely on the weights to balance average characteristics with the community-based ECD sample. Although the method requires that unobserved characteristics do not systematically differ among children in different experimental groups with the same propensity score, we rely on a rich set of predetermined characteristics, including baseline MDAT scores, to generate the weights. The next section presents details of propensity score estimation. Moreover, the method consistently estimates the average treatment effect even if the propensity score equation is misspecified (Robins and Rotnitzky 1995; Imbens and Wooldridge 2009). Under this identifying assumption, we can also compare the ECD Annex experiment to the community-based ECD treatment through a test of the null hypothesis $\beta_2 - \beta_1 = 0$.

Finally, we can compare the ECD Annex control and treatment separately to the community-based ECD groups by disaggregating the *Annex* experiment dummy in equation (B.1) into separate indicators of treatment assignment:

$$y_{ic} = \alpha + \beta_1 community_c + \beta_2 Annex_control_c + \beta_3 Annex_treated_c + \gamma Region2_c + \varepsilon_{ic} \quad (B.2)$$

Estimates of β_2 and β_3 in equation (B.2) help to determine if the Annex setting without training leads to differences with the community-based experiment, or whether the provider training is essential for implementing the new curriculum.

4.5.2 Results

The populations eligible for each experiment live in distinct communities. Villages in the community-based ECD experiment lacked access to structured ECD services prior to the treatment, whereas villages in the ECD Annex experiment already had an ECD facility operating in conjunction with a primary school. These conditions reflect broader differences between households in the two experiments.

To ensure comparability between the two experimental samples, we first estimate the probability that a child is in the ECD Annex sample by running a logit regression on a host of baseline characteristics. These characteristics are the child's age (in months); dummies for female, region, and previous ECD attendance; baseline z-scores in fine motor skills, language and hearing, and height-for-age; mother's schooling (in years); household expenditure per capita and willingness to pay for ECD services as a percentage of this expenditure; the proportion of 17 recommended vaccines received by the child; and an index of mother's mental distress, measured as the proportion

²¹ Section 5 presents details of propensity score estimation.

of 11 mental health issues she reports experiencing “most of the time.” We then exclude children whose propensity score falls outside the common support of the distributions from each experiment, leaving an estimation sample that includes 648 of the 844 children (77%) with an endline MDAT score.²²

In Table A1, we check for balance in baseline fine motor skills and language and hearing using the trimmed sample. MDAT scores differ significantly between experiments when the data are unweighted, for both overall scores on each module and for most subsets of test items (columns 1-4). When weighting observations by the inverse of their propensity score (columns 5-8), no score differs at conventional significance levels.²³ These results give us confidence that reweighting generates an appropriate sample with which to compare experimental treatments.

In Table A2, we present results from inverse propensity score-weighted estimation of the effect of each experimental treatment. Column 1 shows estimates of equation (B.1) for language and hearing skills. The coefficient on the community-based ECD treatment indicator corresponds to a decrease of 0.64 standard deviation in language skills relative to the pure control group (the omitted category), significant at 5%. The effect attenuates somewhat to -0.37 when controlling for baseline score in column 2 but remains significant. These results are striking, given that the pure control group lacks ready access to structured ECD services and the variation between these groups is experimental. The estimates also contrast with the previous estimates of the community-based ECD treatment effect of -0.16 and -0.17 standard deviation, which were not statistically distinguishable from zero (Table 5, columns 3-4).

A clue to this discrepancy lies in the pure control group mean language score of 0.49 (reported halfway down Table 8), which is considerably higher than the unweighted mean of -0.07 (Table 5, column 1). In other words, the reweighting used to estimate equation (B.1) gives greater influence to relatively better-off children, who fare better at home than in community-based ECD. This interpretation also echoes the subgroup analysis of Table 6, in which we found the same pattern among more advantaged children in the community-based ECD experiment. The results also match broader lessons from the literature on the heterogeneous effects of pre-school by socioeconomic status.

Returning to column 1 of Table A2, the coefficient on the pooled ECD Annex group is -0.17, but not statistically significant, indicating no distinguishable difference in outcomes between observationally equivalent children in pure control and ECD Annex communities. However, a test of the difference between the ECD Annex and community-based ECD coefficients, reported at the bottom of the table, shows an advantage of 0.47 standard deviation for the ECD Annex group, significant at 1%. The Lee (2009) bounds for this estimate (reported in brackets) exclude zero, alleviating concerns that differential attrition between experiments drives the result.²⁴ This effect falls to 0.40 standard deviation and remains significant at 10% when controlling for baseline score in column 2. These results suggest that providing ECD services through annexes attached to primary

²² Figure SA2 of the Supplemental Appendix shows the propensity score distributions. The supplemental appendix is available upon request.

²³ The result is not simply a mechanical consequence of reweighting. Although the overall score on each module enters the model for the propensity score, it is not obvious that the procedure would also succeed in balancing the means of each subset of items.

²⁴ The Lee (2009) bounds apply to a pairwise comparison between groups, but continue to weight by the inverse propensity score and adjust for regional stratification as in the regressions. We cannot bound the effect while controlling for the baseline outcome because there must be variation in treatment within covariate cells, whereas normalized MDAT scores are continuous.

schools is more effective at promoting language development than community-based ECD services, at least during our sample period.

In Table A2, columns 3 and 4 disaggregate the ECD Annex group by treatment and control, corresponding to equation (B.2). The ECD Annex treatment and control have nearly identical effects relative to both the pure control group (as shown by their coefficients) and to the community-based ECD treatment (as shown by the estimates reported at the bottom of the table). Bounds for these estimates are also nearly identical. These results suggest that the teacher training provided in the ECD Annex treatment added little value relative to the control, at least by the end of the sample period. Instead, it appears that other features of the ECD Annex environment explain its differences with the community-based ECD program.

Columns 5 to 8 of Table A2 repeat the exercise for fine motor skills. For this outcome, we find no statistically significant differences between the pure control group and the community-based ECD treatment. Nor are there significant differences between the pure control and the ECD Annex sample, either pooled or separately by treatment status. Children in the ECD Annex group do score significantly higher than those in the community-based treatment, however, with magnitudes ranging from 0.35-0.42 standard deviation, though the effects lose significance when controlling for baseline outcome. As with language and hearing, for fine motor skills we again see no discernible benefit to the ECD Annex treatment relative to the control.

5. Conclusion

This paper evaluated two experiments to improve early childhood development services in The Gambia. The first experiment focused on increasing access to ECD services by constructing community-based centers in communities where no structured ECD program existed. The second experiment focused on improving quality of existing ECD centers by training teachers to deliver a new curriculum. We found no evidence that either intervention affected average levels of child development, as measured by tests of fine motor skills or language and hearing. Although the experimental evidence comes from only two of six regions in the country, the same programs were implemented at national scale, making these findings directly relevant for policy.

Exploratory analysis suggests that children from more advantaged households developed language skills more slowly when exposed to the community-based ECD treatment. Parents of these children did not change their home investments, suggesting that these effects are driven by the low quality of community-based ECD centers, not by household responses to treatment.

Additionally, comparisons of observationally similar children across experiments reveal that existing ECD centers increased language skills of the children relative to the community-based alternative, reflecting differences in program quality.

We are only able to document short-term impact estimates, focusing on young children less than two years after treatment assignment. Given the importance of early childhood circumstances for adult outcomes, longer-term measures are required to account fully for the effects of the interventions studied here.

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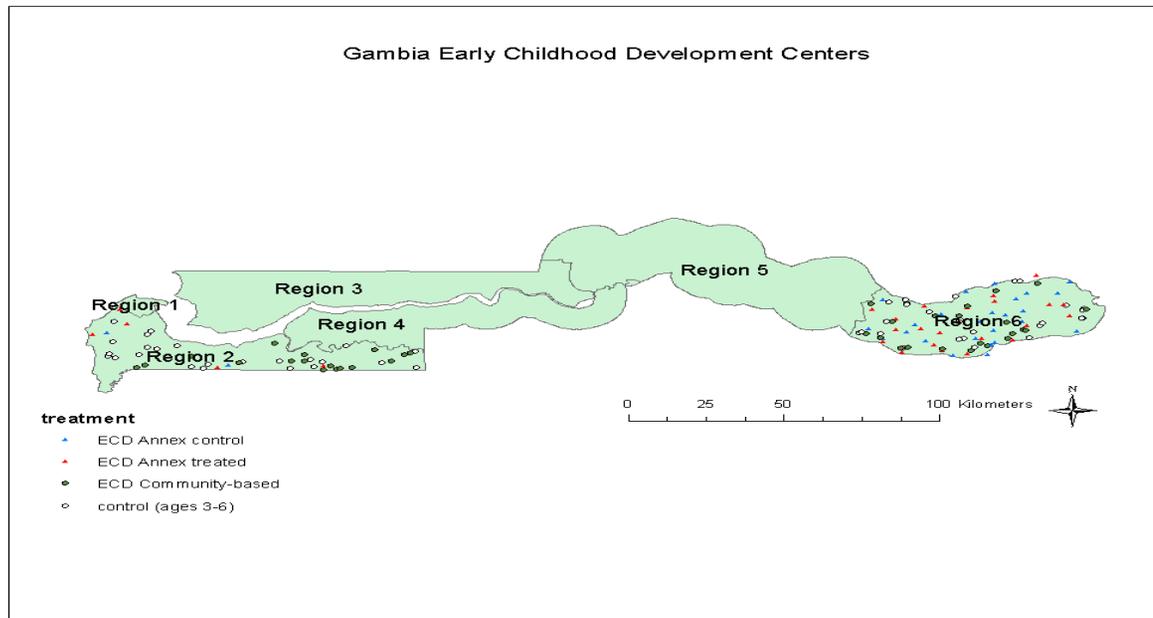
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Figure 1: Sample locations



Source: Authors' elaboration using school location data from Gambia Ministry of Basic and Secondary Education (MoBSE).

Figure 2: Project and research timeline

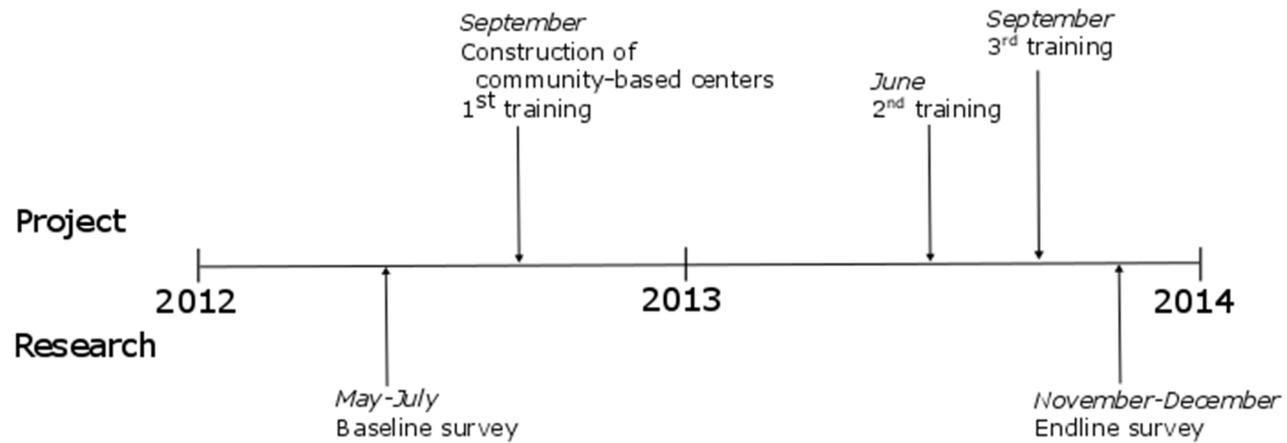


Figure 3(a): Baseline MDAT scores, by age

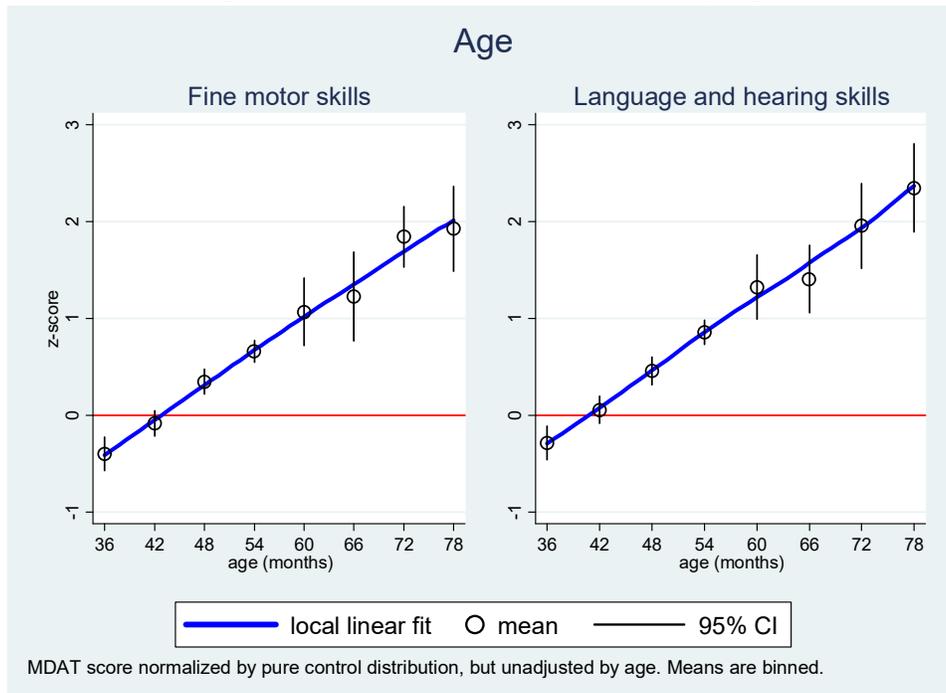
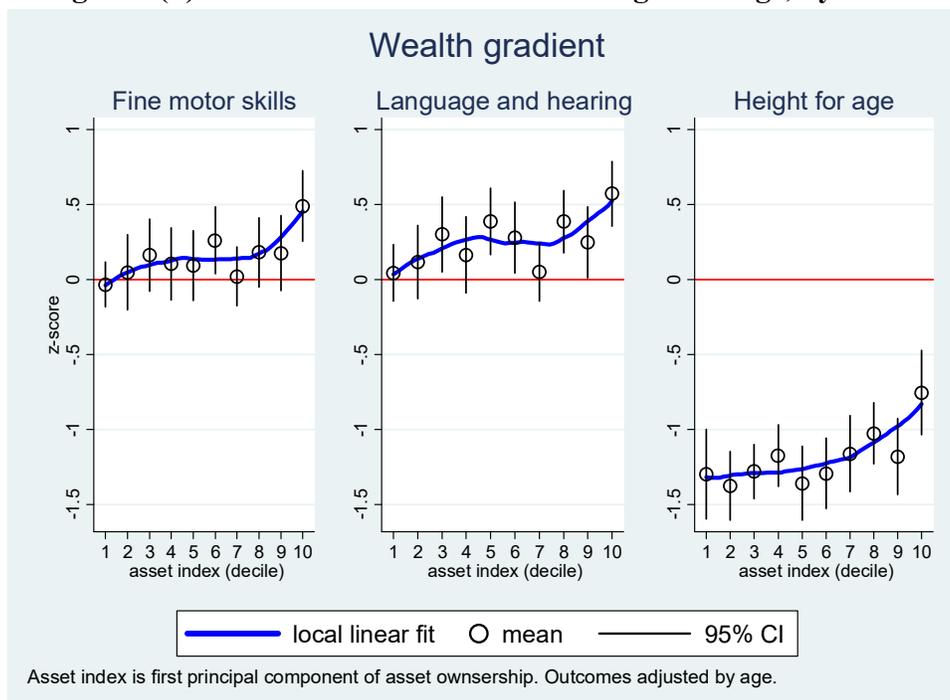


Figure 3(b): Baseline MDAT scores and height-for-age, by wealth



Notes: Figure 3(a) plots unadjusted MDAT scores by child age. Figure 3(b) plots adjusted MDAT scores and height-for-age against deciles of a household wealth index. The wealth index is the first principal component of household asset ownership. Figure 3(b) adjusts MDAT scores for age by regressing a child's raw score on age (in months), age squared, and a female dummy, then dividing the residual by its standard deviation. We run this procedure separately for the fine motor and language and hearing modules. Height-for-age calculated using the international benchmarks defined by the World Health Organization.

Figure 4(a): Outcome distributions, Community-based ECD experiment

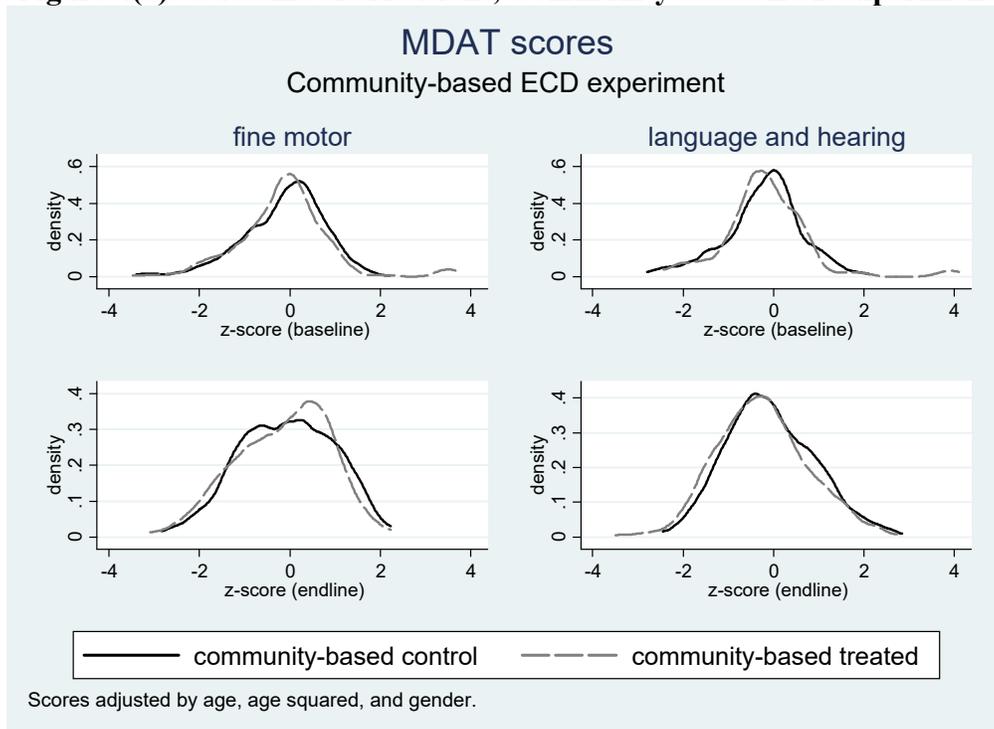
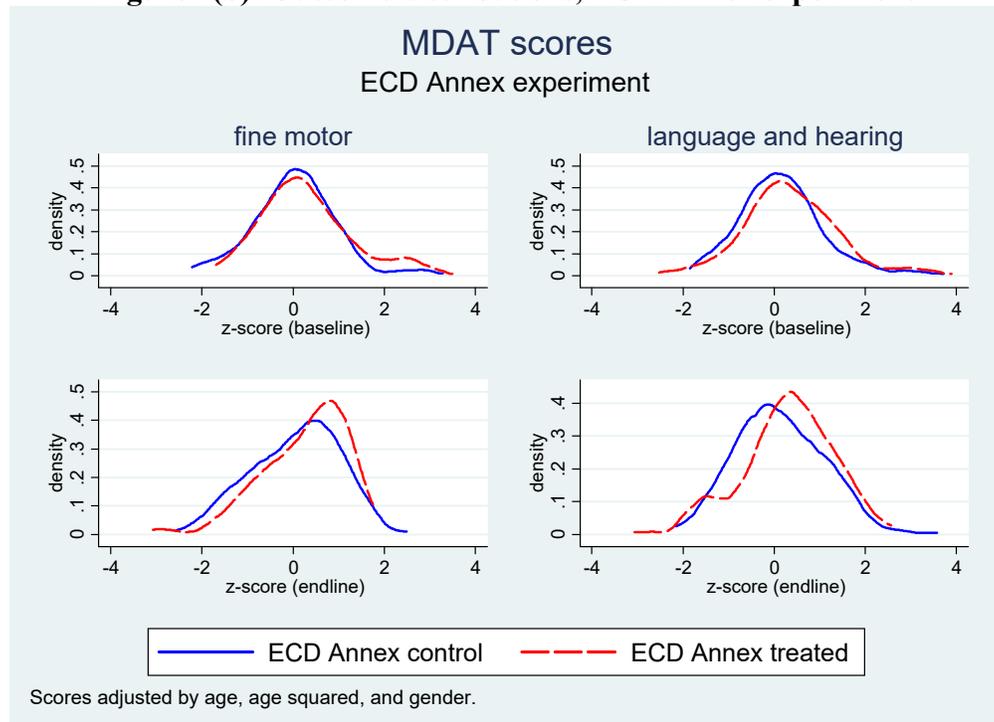


Figure 4(b): Outcome distributions, ECD Annex experiment



Notes: Figure shows kernel density estimates of MDAT scores by survey wave and experimental subgroup. MDAT scores adjusted for age and gender by regressing a child’s raw score on age (in months), age squared, and a female dummy, then dividing the residual by its standard deviation. We run this procedure separately for the fine motor and language and hearing modules and for baseline and endline. Density estimates use Epanechnikov kernel with plug-in bandwidth.

Figure 5(a): Quantile treatment effects, Community-based ECD experiment

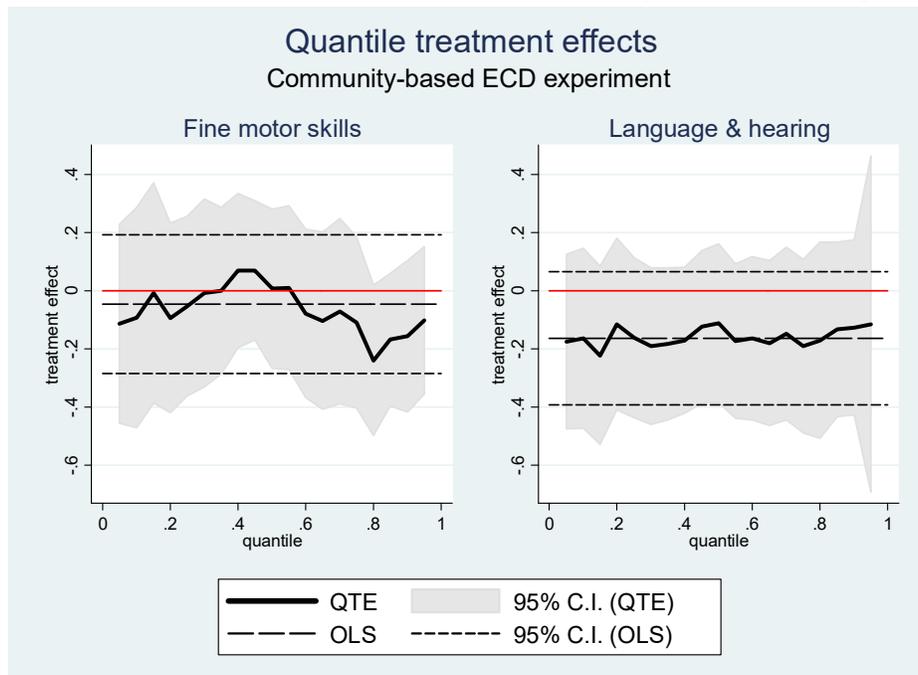
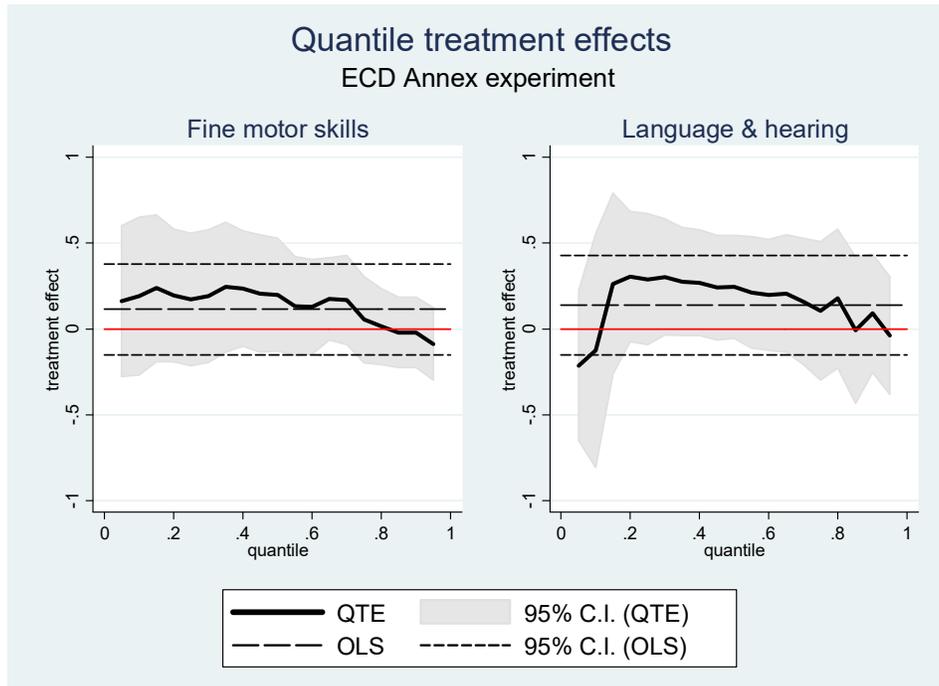


Figure 5(b): Quantile treatment effects, ECD Annex experiment



Notes: Graphs plot the coefficient on treatment status in quantile regressions of equation (1) for the 5th-95th (conditional) quantiles, in increments of 5. Standard errors clustered by community. The thick black line plots estimates of the quantile treatment effects, with 95% confidence interval shaded gray. Graphs also plot the OLS estimate in the long-dashed line (95% confidence interval in short-dashed line). The red line shows the gridline for zero treatment effect.

Table 2: Sample sizes and attrition

	(1)	(2)	(3)	(4)	(5)
Panel A: sample sizes	<u>interview</u>		<u>test score</u>		
	<u>baseline</u>	<u>endline</u>	<u>baseline</u>	<u>Endline</u>	<u>Both</u>
<u>Community-based ECD experiment</u>					
<i>control</i>					
original sample	606	481	319	270	243
added sample		133		75	
<i>treatment</i>					
original sample	441	356	267	226	204
added sample		87		55	
<u>ECD Annex experiment</u>					
<i>control</i>					
original sample	365	322	192	183	170
added sample		44		24	
<i>treatment</i>					
original sample	351	326	182	165	162
added sample		40		23	
Panel B: attrition	<u>control</u>	<u>treatment</u>	<u>difference</u>	<u>p-value</u>	
<u>Community-based ECD experiment</u>					
from endline survey	0.21 (0.03)	0.19 (0.04)	-0.01 (0.05)	0.99	
from endline test	0.24 (0.04)	0.24 (0.04)	0.00 (0.06)	0.82	
number of communities	51	40			
<u>ECD Annex experiment</u>					
from endline survey	0.12 (0.02)	0.07 (0.01)	-0.05 (0.03)	0.11	
from endline test	0.12 (0.03)	0.11 (0.02)	-0.01 (0.03)	0.83	
number of communities	27	26			

Panel A shows sample sizes by survey wave and treatment status. Original sample refers to those present at baseline. Added sample refers to new subjects added at endline who were not present at baseline. Interview refers to completed interview. Test score refers to completed MDAT test for fine motor skills and language/hearing. Panel B shows attrition rates by treatment group. Attrited from endline is indicator for not being present for endline interview, conditional on being present for baseline interview. Attrited from endline test is indicator for not being present for endline test, conditional on taking baseline test. Standard errors in parentheses, clustered by settlement. p-values adjusted for stratification of treatment by region.

Table 3: Baseline characteristics and balance tests

Experiment Group Variable	<u>Community-based ECD</u>				<u>ECD Annex</u>			
	<u>control</u>	<u>treatment</u>	<u>difference</u>	<u>p-value</u>	<u>Control</u>	<u>treatment</u>	<u>difference</u>	<u>p-value</u>
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
age (months)	46.9 (0.4)	47.6 (0.5)	0.8 (0.7)	0.39	57.2 (1.4)	56.1 (1.3)	-1.1 (1.9)	0.75
female	0.49 (0.03)	0.44 (0.03)	-0.06 (0.04)	0.13	0.47 (0.04)	0.50 (0.03)	0.03 (0.05)	0.61
attend ECD	0.37 (0.05)	0.23 (0.04)	-0.14 (0.06)	0.04	0.88 (0.04)	0.92 (0.03)	0.05 (0.05)	0.47
fine motor skills	-0.12 (0.08)	-0.13 (0.13)	0.00 (0.15)	0.78	0.04 (0.12)	0.29 (0.16)	0.25 (0.20)	0.33
language and hearing skills	-0.22 (0.08)	-0.15 (0.13)	0.07 (0.15)	0.47	0.17 (0.12)	0.36 (0.16)	0.19 (0.19)	0.45
height-for-age	-1.35 (0.08)	-1.55 (0.09)	-0.20 (0.12)	0.11	-1.37 (0.12)	-1.28 (0.21)	0.09 (0.24)	0.64
household size	7.7 (0.3)	8.2 (0.3)	0.5 (0.4)	0.27	7.6 (0.3)	8.0 (0.4)	0.4 (0.5)	0.48
mother's schooling	1.8 (0.3)	1.7 (0.3)	0.0 (0.4)	0.67	1.7 (0.2)	2.2 (0.4)	0.5 (0.4)	0.60
household head schooling	2.2 (0.3)	1.8 (0.3)	-0.4 (0.5)	0.88	2.4 (0.4)	2.2 (0.4)	-0.2 (0.6)	0.13
household head employed	0.82 (0.03)	0.84 (0.03)	0.03 (0.05)	0.43	0.82 (0.04)	0.73 (0.05)	-0.09 (0.07)	0.10
household head work hours	38.0 (2.5)	37.8 (2.4)	-0.2 (3.4)	0.95	38.3 (2.7)	33.4 (3.3)	-4.9 (4.3)	0.23
household head in agriculture	0.57 (0.04)	0.72 (0.04)	0.15 (0.06)	0.05	0.60 (0.04)	0.55 (0.05)	-0.05 (0.07)	0.86
household expenditure per capita	483.2 (67.9)	523.6 (56.3)	40.4 (87.8)	0.56	519.0 (40.2)	418.4 (29.4)	-100.6 (49.4)	0.06
asset index	-0.04 (0.22)	-0.52 (0.13)	-0.48 (0.25)	0.05	-0.03 (0.28)	0.68 (0.35)	0.72 (0.44)	0.14

ECD willingness to pay (as % of household expenditure)	0.11 (0.02)	0.21 (0.07)	0.10 (0.07)	0.10	0.07 (0.01)	0.18 (0.06)	0.11 (0.06)	0.11
vaccinations (% of 17)	0.55 (0.03)	0.60 (0.02)	0.05 (0.04)	0.21	0.54 (0.03)	0.56 (0.04)	0.02 (0.05)	0.78
child ill	0.13 (0.02)	0.17 (0.02)	0.04 (0.03)	0.26	0.18 (0.03)	0.16 (0.03)	-0.02 (0.04)	0.53
mother mental distress (0-100 scale, where 100 is worst)	44.9 (1.7)	41.8 (2.4)	-3.1 (2.9)	0.22	47.6 (2.9)	45.8 (3.0)	-1.8 (4.1)	0.64
stimulating objects (% of 10)	0.27 (0.01)	0.28 (0.01)	0.00 (0.02)	0.85	0.28 (0.01)	0.30 (0.01)	0.01 (0.02)	0.58
corporal punishment (% use)	0.71 (0.03)	0.66 (0.04)	-0.05 (0.05)	0.38	0.63 (0.05)	0.69 (0.05)	0.07 (0.06)	0.51
Observations	319	267			192	182		

Sample is children with baseline MDAT score. All variables are means from baseline survey. Standard errors in parentheses, clustered by settlement. p-values adjusted for stratification of treatment by region. Fine motor, language and hearing skills are z-scores from MDAT (adjusted scores based on standardized residuals from regression of raw score on child's age, age squared, and female dummy). Height-for-age z-score based on World Health Organization 2007 benchmark. Household expenditure per capita is annual value in USD, winsorized at 1st/99th percentiles. Asset index is first principal component of reported household assets. ECD willingness to pay is stated willingness to pay for early childhood development services as share of household per capital expenditure. Vaccinations is proportion of 17 vaccinations received by child. Mother mental distress is percentage of 11 mental health issues experienced "most of the time" by mother. Stimulating objects is proportion of 10 objects for stimulating play found in home. Corporal punishment is indicator for using corporal punishment as usual form of discipline for severe misbehavior.

Table 4: Implementation measures

Variable	ECD Annex treated (1)	ECD Community-based (2)	difference (3)	p-value (4)
<i>Site characteristics</i>				
registered children (total)	47.8 (4.04)	45.4 (4.64)	0.4 (6.61)	0.95
children present (total)	32.2 (3.25)	38.4 (4.00)	-6.6 (5.65)	0.25
attendance % (total)	0.71 (0.04)	0.88 (0.04)	-0.14 (0.05)	0.01
pupil-teacher ratio	35.7 (4.01)	31.4 (3.97)	-0.4 (5.38)	0.94
<i>Teacher characteristics</i>				
female	0.32 (0.08)	0.38 (0.07)	0.07 (0.11)	0.81
age	33.3 (1.3)	27.0 (0.9)	-6.3 (1.6)	0.00
at least primary	0.64 (0.08)	0.85 (0.04)	0.21 (0.09)	0.02
at least secondary	0.61 (0.07)	0.30 (0.06)	-0.31 (0.09)	0.00
qualified	0.50 (0.11)	0.03 (0.03)	-0.47 (0.12)	0.00
teaching experience	5.7 (1.3)	0.5 (0.2)	-5.1 (1.3)	0.00
Ever taught at ECD	0.54 (0.11)	0.13 (0.05)	-0.41 (0.12)	0.00
ECD experience	1.3 (0.3)	0.3 (0.1)	-1.0 (0.4)	0.01
Citizen of the village	0.16 (0.08)	0.65 (0.07)	0.49 (0.10)	0.00
commute time (minutes)	18.2 (3.4)	15.2 (2.4)	-3.0 (4.2)	0.43
work hours last week	18.5 (2.0)	17.8 (1.9)	-0.6 (2.8)	0.95
absent from class last month	0.42 (0.08)	0.21 (0.08)	-0.21 (0.11)	0.03
named theme from syllabus as this week's topic	0.57 (0.06)	0.43 (0.06)	-0.14 (0.09)	0.21
follows syllabus fully	0.43 (0.06)	0.37 (0.06)	-0.07 (0.08)	0.66
<u>To what extent are children...? Very much</u>				
...asking questions?	0.54 (0.11)	0.72 (0.08)	0.19 (0.13)	0.17
...taking initiative?	0.54 (0.10)	0.76 (0.08)	0.22 (0.12)	0.08
...interacting among themselves?	0.82 (0.08)	0.86 (0.06)	0.04 (0.10)	0.72
...listening and responding to teacher?	0.75 (0.09)	0.90 (0.06)	0.15 (0.11)	0.17
...interacting with toys and books?	0.89 (0.06)	0.90 (0.06)	0.01 (0.08)	0.94
<u>Observations</u>				
Teachers	44	60		
Sites	25	38		

Table shows results of monitoring survey of ECD providers in treatment sites, 2013. Teaching experience measured in years.

Table 5: Endline outcomes

Experiment Group	Community-based ECD				ECD Annex			
	control	treatment	treatment effect		control	treatment	treatment effect	
Variable	(1)	(2)	without baseline outcome	with baseline outcome	(5)	(6)	without baseline outcome	with baseline outcome
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
<u>Language and hearing</u>								
overall score (z)	-0.07 (0.08)	-0.25 (0.10)	-0.16 (0.12)	-0.17 (0.12)	0.12 (0.10)	0.30 (0.12)	0.14 (0.15)	0.11 (0.15)
knows own name & its letters (% of 4)	0.99 (0.01)	0.98 (0.01)	-0.01 (0.01)	-0.01 (0.01)	1.00 (0.01)	1.00 (0.01)	0.00 (0.01)	0.00 (0.01)
speaks in clear sentences	0.99 (0.01)	0.97 (0.02)	-0.02 (0.02)	-0.04 (0.02)	0.99 (0.01)	0.99 (0.01)	0.00 (0.01)	0.00 (0.01)
counting (% of 3)	0.39 (0.02)	0.43 (0.03)	0.04 (0.04)	0.04 (0.04)	0.49 (0.02)	0.45 (0.02)	-0.04 (0.03)	-0.09** (0.03)
name colors (% of 4)	0.16 (0.01)	0.19 (0.02)	0.02 (0.03)	0.01 (0.03)	0.24 (0.01)	0.25 (0.02)	0.02 (0.02)	-0.01 (0.03)
play with blocks (% of 3)	0.38 (0.02)	0.37 (0.02)	0.00 (0.03)	N/A	0.45 (0.02)	0.43 (0.02)	-0.02 (0.03)	N/A
open books (% of 3)	0.18 (0.01)	0.17 (0.02)	0.00 (0.02)	N/A	0.23 (0.02)	0.19 (0.02)	-0.04 (0.03)	N/A
<u>Fine motor skills</u>								
overall score (z)	-0.07 (0.10)	-0.15 (0.12)	-0.05 (0.12)	-0.04 (0.14)	0.06 (0.10)	0.24 (0.12)	0.11 (0.14)	0.16 (0.14)
play with blocks (% of 6)	0.38 (0.02)	0.40 (0.02)	0.03 (0.03)	-0.01 (0.03)	0.42 (0.03)	0.40 (0.02)	-0.02 (0.03)	-0.02 (0.03)
draw lines & shapes (% of 6)	0.41 (0.02)	0.45 (0.02)	0.05 (0.03)	0.03 (0.03)	0.45 (0.03)	0.43 (0.02)	-0.02 (0.03)	-0.04 (0.03)
order rows of items (% of 2)	0.29 (0.02)	0.29 (0.02)	0.00 (0.03)	-0.01 (0.03)	0.39 (0.03)	0.34 (0.02)	-0.06 (0.04)	-0.08* (0.04)
fold paper	0.62 (0.04)	0.62 (0.05)	0.01 (0.06)	N/A	0.68 (0.06)	0.76 (0.05)	0.07 (0.08)	N/A
play with blocks (language and motors skills combined, % of 6)	0.38 (0.02)	0.39 (0.02)	0.01 (0.03)	N/A	0.43 (0.02)	0.41 (0.02)	-0.02 (0.03)	N/A
<u>Observations</u>								
children	345	281			207	188		
sites	50	40			27	26		

* significant at 10%, ** significant at 5%, *** significant at 1%. All variables are means from MDAT endline survey. Standard errors in parentheses, clustered by settlement. Treatment effects control for stratification of treatment assignment by region. Estimates in columns (4) and (8) augment regression with control for baseline outcome. z-scores are standardized residuals from regression of raw score on child's age, age squared, and female dummy. Other variables are subsets of items on MDAT test, measured as percent of items completed correctly. Speaks in clear sentences is just one item, while other categories have number of items indicated.

Table 6(a): Treatment effect heterogeneity, Community-based ECD experiment

Outcome	Fine motor skills						Language and hearing					
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
<u>Treatment interacted with:</u>												
male	0.03 (0.11)						-0.10 (0.11)					
female	-0.09 (0.11)						-0.28* (0.11)					
assets below median		-0.01 (0.12)						-0.13 (0.12)				
assets above median		-0.12 (0.13)						-0.40** (0.13)				
mother didn't attend school			0.05 (0.09)						-0.12 (0.09)			
mother attended school			-0.32 (0.17)						-0.40* (0.17)			
stimulating objects below median				0.14 (0.10)						-0.04 (0.10)		
stimulating objects above median				-0.24* (0.12)						-0.36** (0.12)		
child not sick last 3 days					-0.05 (0.08)						-0.20* (0.08)	
child sick last 3 days					0.11 (0.22)						-0.01 (0.23)	
mother mental distress below median						-0.04 (0.11)						-0.24* (0.11)
mother mental distress above median						-0.09 (0.14)						-0.28* (0.14)
N			622						618			
H1: low SES treatment effect=0			0.59						0.59			
H2: high SES treatment effect=0			0.17						0.05			
H3: low SES = high SES			0.02						0.01			

* significant at 10%, ** significant at 5%, *** significant at 1%. Fine motor skills and language and hearing outcomes are MDAT z-scores, adjusted for age, age squared, and gender. Each column shows results of a separate regression. All regressions include control for Region 2 and dummies for all subgroups listed in table (constant omitted so that all subgroup coefficients and interactions identified). p-values reported at bottom from tests across equations, using joint variance-covariance matrix of all coefficients. "Low SES treatment effect=0" reports p-value of joint test on interactions between treatment and indicators for assets below median, mother didn't attend school, and stimulating objects below median. "High SES treatment effect=0" reports p-value of joint test on interactions between treatment and indicators for assets above median, mother attended school, and stimulating objects above median. "Low SES=High SES" reports p-value of joint test of equality between low-SES interactions and high-SES interactions. Household assets based on first principal component of asset ownership. ECD demand is stated willingness to pay for ECD services as percentage of household per capita expenditure. Mother mental distress is proportion of responses to mental health questionnaire answering "3," which is "most of the time" for mental health issue. Stimulating objects in home based on proportion of 10 objects owned. Standard errors in parenthesis, clustered by settlement.

Table 6(b): Treatment effect heterogeneity, ECD Annex experiment

Outcome	Fine motor skills					Language and hearing						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
<u>Treatment interacted with:</u>												
male	0.24 (0.13)						0.26 (0.14)					
female	-0.03 (0.14)						-0.01 (0.15)					
assets below median		0.09 (0.15)						0.07 (0.16)				
assets above median		0.21 (0.13)						0.15 (0.15)				
mother didn't attend school			0.08 (0.11)						0.10 (0.12)			
mother attended school			0.23 (0.19)						0.22 (0.20)			
stimulating objects below median				0.05 (0.13)						-0.08 (0.14)		
stimulating objects above median				0.18 (0.13)						0.34* (0.15)		
child not sick last 3 days					0.16 (0.10)						0.17 (0.11)	
child sick last 3 days					-0.10 (0.24)						-0.08 (0.26)	
mother mental distress below median						-0.04 (0.13)						0.00 (0.15)
mother mental distress above median						0.39** (0.15)						0.26 (0.16)
N			392						388			
H1: low SES treatment effect=0			0.95						0.39			
H2: high SES treatment effect=0			0.66						0.35			
H3: low SES = high SES			0.31						0.16			

* significant at 10%, ** significant at 5%, *** significant at 1%. Fine motor skills and language and hearing outcomes are MDAT z-scores, adjusted for age, age squared, and gender. Each column shows results of a separate regression. All regressions include control for Region 2 and dummies for all subgroups listed in table (constant omitted so that all subgroup coefficients and interactions identified). p-values reported at bottom from tests across equations, using joint variance-covariance matrix of all coefficients. "Low SES treatment effect=0" reports p-value of joint test on interactions between treatment and indicators for assets below median, mother didn't attend school, and stimulating objects below median. "High SES treatment effect=0" reports p-value of joint test on interactions between treatment and indicators for assets above median, mother attended school, and stimulating objects above median. "Low SES=High SES" reports p-value of joint test of equality between low-SES interactions and high-SES interactions. Household assets based on first principal component of asset ownership. ECD demand is stated willingness to pay for ECD services as percentage of household per capita expenditure. Mother mental distress is proportion of responses to mental health questionnaire answering "3," which is "most of the time" for mental health issue. Stimulating objects in home based on proportion of 10 objects owned. Standard errors in parenthesis, clustered by settlement.

Table 7: Home investments (Time)

Experiment	Outcome: mother play time with child (minutes/day)													
	Community-based ECD							ECD Annex						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)
Treatment	5.6 (19.4)							46.2 (27.9)						
<u>Treatment interacted with:</u>														
male		-8.1 (15.6)							54.6** (17.8)					
female		20.0 (16.9)							36.9* (18.6)					
assets below median			12.0 (17.5)							25.8 (20.8)				
assets above median			13.5 (19.0)							56.5** (18.3)				
mother didn't attend school				1.4 (12.9)							56.7*** (14.6)			
mother attended school				18.2 (25.4)							9.3 (27.0)			
stimulating objects below median					-2.0 (15.2)							30.3 (18.8)		
stimulating objects above median					14.0 (17.5)							61.0*** (18.2)		
child not sick last 3 days						10.7 (12.3)							36.2** (14.0)	
child sick last 3 days						-34.8 (32.0)							104.1** (33.4)	
mother mental distress below median							3.3 (16.5)							43.2* (18.2)
mother mental distress above median							25.6 (20.5)							41.0 (21.1)
N	1,045	1,045	1,045	1,045	1,045	1,045	1,045	728	728	728	728	728	728	728
Control group mean	184.5							162.4						
H1: low SES treatment effect=0				0.89							0.02			
H2: high SES treatment effect=0				0.92							0.04			
H3: low SES = high SES				0.50							0.78			

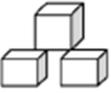
* significant at 10%, ** significant at 5%, *** significant at 1%. Outcome is mother's play time with child (minutes/day). Each column shows results of a separate regression. All regressions include control for Region 2. Regressions with interacted treatment variables include dummies for main effect of all subgroups listed in table (constant omitted so that all subgroup coefficients and interactions identified). p-values reported at bottom from tests across equations, using joint variance-covariance matrix of all coefficients. "Low SES treatment effect=0" reports p-value of joint test on interactions between treatment and indicators for assets below median, mother didn't attend school, and stimulating objects below median. "High SES treatment effect=0" reports p-value of joint test on interactions between treatment and indicators for assets above median, mother attended school, and stimulating objects above median. "Low SES=High SES" reports p-value of joint test of equality between low-SES interactions and high-SES interactions. Household assets based on first principal component of asset ownership. ECD demand is stated willingness to pay for ECD services as percentage of household per capita expenditure. Mother mental distress is proportion of responses to mental health questionnaire answering "3," which is "most of the time" for mental health issue. Stimulating objects in home based on proportion of 10 objects owned. Standard errors in parenthesis, clustered by settlement.

Table 8: Home investments (Stimulation Items)

Experiment	Outcome: stimulating items in household (proportion of 11)													
	Community-based ECD							ECD Annex						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)
Treatment	0.03 (0.04)							-0.02 (0.04)						
<u>Treatment interacted with:</u>														
male		0.05** (0.02)							-0.02 (0.03)					
female		0.00 (0.02)							-0.04 (0.03)					
assets below median			0.05* (0.02)							-0.04 (0.03)				
assets above median			-0.01 (0.02)							-0.04 (0.03)				
mother didn't attend school				0.03 (0.02)							-0.02 (0.02)			
mother attended school				0.01 (0.03)							-0.05 (0.04)			
stimulating objects below median					0.03 (0.02)							-0.02 (0.03)		
stimulating objects above median					0.03 (0.02)							-0.04 (0.03)		
child not sick last 3 days						0.02 (0.02)							-0.03 (0.02)	
child sick last 3 days						0.05 (0.04)							0.00 (0.05)	
mother mental distress below median							0.03 (0.02)							-0.04 (0.03)
mother mental distress above median							0.00 (0.03)							-0.04 (0.03)
N	1,055	1,055	1,055	1,055	1,055	1,055	1,055	730	730	730	730	730	730	730
Control group mean	0.30							0.36						
H1: low SES treatment effect=0					0.61						0.75			
H2: high SES treatment effect=0					0.57						0.74			
H3: low SES = high SES					0.35						0.66			

* significant at 10%, ** significant at 5%, *** significant at 1%. Outcome is stimulating objects found in home (proportion of list of 11). Each column shows results of a separate regression. All regressions include control for Region 2. Regressions with interacted treatment variables include dummies for main effect of all subgroups listed in table (constant omitted so that all subgroup coefficients and interactions identified). p-values reported at bottom from tests across equations, using joint variance-covariance matrix of all coefficients. "Low SES treatment effect=0" reports p-value of joint test on interactions between treatment and indicators for assets below median, mother didn't attend school, and stimulating objects below median. "High SES treatment effect=0" reports p-value of joint test on interactions between treatment and indicators for assets above median, mother attended school, and stimulating objects above median. "Low SES=High SES" reports p-value of joint test of equality between low-SES interactions and high-SES interactions. Household assets based on first principal component of asset ownership. ECD demand is stated willingness to pay for ECD services as percentage of household per capita expenditure. Mother mental distress is proportion of responses to mental health questionnaire answering "3," which is "most of the time" for mental health issue. Stimulating objects in home based on proportion of 10 objects owned. Standard errors in parenthesis, clustered by settlement.

Appendix A.1: MDAT, baseline

MDAT FINE MOTOR 36-59 MONTHS: ITEMS 18-41				
<p>23. USE 12 blocks. Make tower of 6 BLOCKS and then put 6 blocks in front of child and say "Make a tower just like mine." Allow 2 trials.</p>  <p>YES.....1 NO.....0</p>	<p>22. Child can stack 4 blocks. SCORE FROM ITEM 23</p>  <p>YES.....1 NO.....0</p>	<p>20. Child can stack 2 blocks. SCORE FROM ITEM 23</p>  <p>YES.....1 NO.....0</p>	<p>30. USE 6 BLOCKS. Child can make bridge of 3 blocks</p>  <p>YES.....1 NO.....0</p>	<p>35. USE 12 BLOCKS. Child can make bridge of 6 blocks</p>  <p>YES.....1 NO.....0</p>
<p>36. USE 12 BLOCKS. Child can make stairs with 6 blocks</p>  <p>YES.....1 NO.....0</p>	<p>24. Show child how fingers go into peg holes. Put in 2 pegs, explaining what you are doing. Remove pegs and ask child to put them all in. Child puts pegs into board in up <= 30 secs.</p>  <p>YES.....1 NO.....0</p>	<p>21. Child puts pegs into board in up to 2 minutes SCORE FROM ITEM 24</p>  <p>YES.....1 NO.....0</p>	<p>25. Screws and unscrews lid on jar (after demonstration).</p>  <p>YES.....1 NO.....0</p>	<p>26. Threads 6 beads (demonstrate with 2 beads).</p>  <p>YES.....1 NO.....0</p>

MDAT LANGUAGE AND HEARING 36-59 MONTHS: ITEMS 16-43

<p>20. MARK 'YES' IF CHILD HAS TOLD YOU HIS/HER NAME. If not, ask child to tell you his or her first name.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>18. Ask child to point to eyes, ears, nose, mouth, hands. SCORE YES IF CHILD POINTS TO >=2 PARTS.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>23. Remove 12 objects. Touch or point to each object, saying "Tell me what this is," or "What is this called?" Mark YES if child can NAME 10 objects</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>19. Child can NAME 5 objects in the basket. SCORE FROM THE ADMINISTRATION OF 23.</p> <p>YES.....1 NO.....0</p>	<p>22. With the 12 objects out, ask child, "Where's the ____?" DO NOT point, touch or look at the objects. Score YES if child can IDENTIFY 10 objects.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>
<p>16. Child can IDENTIFY (point to or give you) 5 or more objects you name. SCORE FROM THE ADMINISTRATION OF 22.</p> <p>YES.....1 NO.....0</p>	<p>21. USE cup, pencil, matchbox, car. Without gesturing, ask "Which one is for drinking?" "Which one is for writing?" "Which one is for fighting fire?" "Which one is for driving?" Score YES if child knows 3 or more.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>26. USE bicycle, spoon, broom, bottle. Without naming, POINT to each object and say, "What do you do with THIS?" Score YES if child correctly uses action words (verbs) for 3 or more.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>24. Say, "Tell me 3 foods you like to eat." Be sure to prompt if child says "tea." Ask for 3 animals if child says 2 or fewer foods. Score YES if >=3 foods OR animals.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>28. Say, "What do you do when you are ...hungry/tired/cold?" Score YES if child can answer 2 of 3.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>
<p>29. Say, "Which goes faster, a car or a bicycle?" Score YES if child says car.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>30. Tell child, "Repeat what I say. Say 'Pa.' Say 'Pa-Chi.' Say 'Pa-Chi-Tu.' Say 'Pa-Chi-Tu-Go.'" Score YES if child repeats all stages correctly.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>25. Say, "Listen carefully to me and do EXACTLY what I say. Ready? OK. I want you to stand up, then clap your hands and then turn around."</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>31. Use cup and counter. Ask child to put counter on top of, under, next to, in front of and behind cup. Must do 3.</p> <p align="center"></p> <p>YES.....1 NO.....0</p>	<p>32. OPPOSITES. "I want you to help me finish some sentences. If a boy is small, a man is... If the sun comes up in the day, the moon comes up ... A baby is young, and a grandma is <u>old</u>. Brothers and fathers are boys, mothers and sisters are <u>girls</u>. You cry when you are sad, but you smile when you are <u>happy</u>. Score YES if 2 or more of these.</p> <p>YES.....1 NO.....0</p>

Appendix A.2: MDAT, endline

ENDLINE MDAT FINE MOTOR: 53-76 MONTHS OF AGE v.5				
<p>34. Can copy a square. CAN TRY 2X</p>  <p>YES.....1</p>  <p>NO.....0</p>	<p>32. Can copy a circle. CAN TRY 2X</p>  <p>YES.....1</p>  <p>NO.....0</p>	<p>43. Child is able to color within the lines of a square. Child should shade in the majority of the square. SCORE YES if fewer than 5 marks outside the lines.</p> <p>YES.....1</p> <p>NO.....0</p>	<p>42. Child is able to fold a piece of paper neatly into quarters. Demonstrate, explaining what you are doing. CAN TRY 2X. Score YES if child folds fairly evenly, with little mismatching (less than 1 inch all around).</p> <p>YES.....1</p> <p>NO.....0</p>	<p>41. Can copy 5 letters. Write E C A M J H and ask child to copy each letter. Score YES for letters with all elements, even if upside down or backwards.</p> <p>YES.....1</p> <p>NO.....0</p>
<p>40. Can copy 3 letters. Write E C A M J H and ask child to copy each letter. Credit given for letters with all elements, even if upside down or backwards. SCORE FROM ITEM 41.</p> <p>YES.....1</p> <p>NO.....0</p>	<p>39. Can copy 1 letter. Write E C A M J H and ask child to copy each letter. Credit given for letters with all elements, even if upside down or backwards. SCORE FROM ITEM 41.</p> <p>YES.....1</p> <p>NO.....0</p>			

ENDLINE MDAT LANGUAGE AND HEARING 53-76 MONTHS OF AGE v.5				
<p>46. Use 12 blocks. Put them in a row. Ask child to hand you <u>one</u> block.</p> <p>YES.....1 NO.....0</p>	<p>47. Child can hand you <u>three</u> blocks.</p> <p>YES.....1 NO.....0</p>	<p>48. Child can hand you <u>five</u> blocks.</p> <p>YES.....1 NO.....0</p>	<p>37. USE PAGE WITH 4 COLORED CIRCLES. Point to the RED circle and ask, "What is the name of this color?" Score YES if child says RED.</p> <p>YES.....1 NO.....0</p>	<p>38. USE PAGE WITH 4 COLORED CIRCLES. Point to the BLUE circle and ask, "What is the name of this color?" Score YES if child says BLUE.</p> <p>YES.....1 NO.....0</p>
<p>39. USE PAGE WITH 4 COLORED CIRCLES. Point to the YELLOW circle and ask, "What is the name of this color?" Score YES if child says BLACK/YELLOW.</p> <p>YES.....1 NO.....0</p>	<p>40. USE PAGE WITH 4 COLORED CIRCLES. Point to the GREEN circle and ask, "What is the name of this color?" Score YES if child says GREEN.</p> <p>YES.....1 NO.....0</p>	<p>49. Present the book to child (sideways). Say, "Here's a fun book. Let's read the book together. Show me how we read the book (OR SAY "SHOW ME HOW YOU READ A BOOK"). Score YES if Child turns book to right orientation.</p> <p>YES.....1 NO.....0</p>	<p>50. Continued from item 49. Say, "Let's read this book together. Show me THE PAGE where we can start reading the book." Score YES if Child turns and points to first page where text is expected to begin.</p> <p>YES.....1 NO.....0</p>	<p>51. Continued from item 50. Say, "Great! Show me where we start reading. Put your finger on the first word where we start reading this story." Score YES if child puts finger on upper left word.</p> <p>YES.....1 NO.....0</p>
<p>*****CHECK***** HAVE THE FOLLOWING ITEMS ADMINISTERED AT THE START OF TEST (AFTER INTRODUCTIONS) BEEN SCORED?</p>	<p>20. MARK 'YES' IF CHILD HAS TOLD YOU HIS/HER NAME. If not, ask child to tell you his or her first name.</p> <p>YES.....1 NO.....0</p>	<p>36. MARK 'YES' IF CHILD HAS TOLD YOU HOW OLD THEY ARE. If not, ask child to tell you his or her age. Verify.</p> <p>YES.....1 NO.....0</p>	<p>44. Ask the child the name of the village where s/he currently lives. Score YES if child answers correctly. Be sure to verify.</p> <p>YES.....1 NO.....0</p>	<p>17. MARK 'YES' IF YOU HEARD CHILD speak clearly in sentences (more than just 2 words put together). IF UNCERTAIN, ask Mom if child uses sentences, even if not grammatically correct, e.g., "Mommy give cup," "Give me ball."</p> <p></p> <p>YES.....1 NO.....0</p>

Table A1: Between-experiments baseline balance tests

	<u>unweighted</u>				<u>weighted</u>			
	<u>community-based</u>	<u>ECD Annex</u>	<u>(1) vs. (2)</u>	<u>p-value</u>	<u>community-based</u>	<u>ECD Annex</u>	<u>(1) vs. (2)</u>	<u>p-value</u>
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
<u>Language and hearing</u>								
overall score (z)	-0.19 (0.07)	0.27 (0.10)	0.58 (0.12)	0.00	0.33 (0.20)	0.09 (0.11)	-0.24 (0.23)	0.30
knows own name & its letters (% of 4)	0.13 (0.01)	0.19 (0.01)	0.06 (0.02)	0.00	0.18 (0.04)	0.17 (0.01)	-0.01 (0.04)	0.87
speaks in clear sentences	0.93 (0.02)	0.96 (0.01)	0.02 (0.02)	0.26	0.92 (0.03)	0.95 (0.02)	0.03 (0.03)	0.36
counting (% of 3)	0.09 (0.01)	0.25 (0.02)	0.17 (0.02)	0.00	0.23 (0.05)	0.19 (0.02)	-0.03 (0.06)	0.55
name colors (% of 4)	0.04 (0.01)	0.10 (0.02)	0.06 (0.02)	0.00	0.13 (0.04)	0.07 (0.01)	-0.07 (0.05)	0.15
<u>Fine motor skills</u>								
overall score (z)	-0.15 (0.08)	0.16 (0.10)	0.46 (0.13)	0.00	0.26 (0.22)	0.03 (0.11)	-0.23 (0.24)	0.33
play with blocks (% of 6)	0.35 (0.02)	0.36 (0.02)	0.03 (0.02)	0.11	0.41 (0.03)	0.36 (0.02)	-0.05 (0.04)	0.16
draw lines & shapes (% of 6)	0.35 (0.02)	0.40 (0.02)	0.07 (0.03)	0.01	0.44 (0.04)	0.39 (0.02)	-0.05 (0.05)	0.28
order rows of items (% of 2)	0.09 (0.01)	0.12 (0.02)	0.06 (0.02)	0.00	0.13 (0.06)	0.11 (0.02)	-0.02 (0.06)	0.67
<u>Observations</u>								
Children	328	320			328	320		
Sites	55	50			55	50		

All variables are means from baseline survey. Drops observations outside common support of propensity score distribution. Columns (5)-(8) weighted by inverse propensity score where indicated. Standard errors in parentheses, clustered by settlement. p-values obtained from regression of characteristic on community-based treatment and Region 2 dummy in order to adjust for stratification by region. Fine motor, language and hearing skills are z-scores from MDAT. Adjusted scores are standardized residuals from regression of raw score on child's age, age squared, and female dummy. Other variables are subsets of items on MDAT test, measured as percent of items completed correctly. Speaks in clear sentences is just one item, while other categories have number of items indicated

Table A2: Endline outcomes, combined experimental groups

MDAT module	Language and hearing				Fine motor			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Community-based ECD	-0.64 (0.28)**	-0.37 (0.18)**	-0.64 (0.28)**	-0.36 (0.18)**	-0.33 (0.25)	-0.14 (0.21)	-0.33 (0.25)	-0.14 (0.21)
ECD Annex	-0.17 (0.29)	0.04 (0.22)			0.06 (0.21)	0.20 (0.16)		
ECD Annex control			-0.15 (0.35)	0.13 (0.31)			0.09 (0.25)	0.26 (0.22)
ECD Annex treatment			-0.19 (0.27)	-0.07 (0.20)			0.02 (0.21)	0.12 (0.15)
<u>Observations</u>								
Children	639	572	639	572	644	590	644	590
Sites	132	127	132	127	132	128	132	128
pure control group mean	0.49 (1.14)	0.49 (1.14)	0.49 (1.14)	0.49 (1.14)	0.23 (0.99)	0.23 (0.99)	0.23 (0.99)	0.23 (0.99)
Includes baseline outcome		x		x		x		x
<u>Difference from Community-based ECD</u>								
ECD Annex	0.47*** (0.18) [0.22,0.74]	0.40* (0.21)			0.39** (0.19) [0.16,0.66]	0.34 (0.21)		
ECD Annex control			0.49* (0.26) [0.15,0.74]	0.49 (0.31)			0.42* (0.23) [0.42,0.76]	0.40 (0.25)
ECD Annex treatment			0.45*** (0.17) [0.15,0.68]	0.29* (0.17)			0.35* (0.19) [0.11,0.58]	0.26 (0.19)

* significant at 10%, ** significant at 5%, *** significant at 1%. Pure control group (i.e., control group from community-based ECD experiment) is omitted category. All regressions include region 2 dummy to adjust for stratification of treatment assignment and weight by inverse propensity score. Sample drops observations outside common support of propensity score distribution. Standard errors in parentheses, clustered by settlement. Outcomes are adjusted z-scores from MDAT modules for language and hearing and fine motor skills. Adjusted scores are standardized residuals from regression of raw score on child's age, age squared, and female dummy. Regressions include baseline outcome where indicated. Differences with community-based ECD reported at bottom of table based on tests of indicated coefficient with community-based ECD. Lee bounds reported in brackets, based on pairwise comparison, but still reweighting by inverse propensity score and stratifying by region. Propensity score obtained from logit model of membership in ECD Annex sample regressed on baseline characteristics. Included baseline characteristics: age (exact based on DOB), female, Region 2, ECD attendance, fine motor skills (age-adjusted z-score), language and hearing (age-adjusted z-score), height-for-age, household size, mother's years of schooling, household expenditure per capita (winsorized at 1st/99th percentiles), willingness to pay for ECD as % of household expenditure per capita, % of vaccines received, mother mental distress (% of items reported as experiencing "most of the time"). Missing values imputed to zero, with dummies for imputed value included as additional covariates in regression.